# Endemic infection can shape exposure to novel pathogens: Pathogen co-occurrence networks in the Serengeti lions

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#### Abstract :

Pathogens are embedded in a complex network of microparasites that can collectively or individually alter disease dynamics and outcomes. Endemic pathogens that infect an individual in the first years of life, for example, can either facilitate or compete with subsequent pathogens thereby exacerbating or ameliorating morbidity and mortality. Pathogen associations are ubiquitous but poorly understood, particularly in wild populations. We report here on 10 years of serological and molecular data in African lions, leveraging comprehensive demographic and behavioural data to test if endemic pathogens shape subsequent infection by epidemic pathogens. We combine network and community ecology approaches to assess broad network structure and characterise associations between pathogens across spatial and temporal scales. We found significant non-random structure in the lion-pathogen co-occurrence network and identified both positive and negative associations underlying pathogen co-occurrence networks.

**Keywords** : Babesia, calicivirus, canine distemper virus, co-infection, community assembly, coronavirus, feline immunodeficiency virus, parvovirus

### 65 Introduction

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Identifying and determining the nature of interactions between multiple pathogens is increasingly 67 considered critical to understanding infectious disease dynamics (e.g., Pedersen & Fenton 2007; 68 69 Graham 2008; Telfer et al. 2010; Johnson et al. 2015; Gorsich et al. 2018). Individuals are often 70 co-infected by a diverse infra-community of pathogens, and interactions between pathogens can both alter infection patterns (Cattadori et al. 2008; Lass et al. 2013; Susi et al. 2015) and 71 72 influence disease outcomes (Moss et al. 2008; Munson et al. 2008; Knowles 2011; Wejse et al. 2015). Pathogens infecting individuals in the first years of life may impact infection by 73 74 subsequent pathogens (Fenton 2008; Randall et al. 2013; Rynkiewicz et al. 2015; Aivelo & 75 Norberg 2018; Budischak et al. 2018). For example, endemic pathogens that compete for the 76 same resources as epidemic pathogens and can reduce the likelihood of infection (Randall et al. 2013) or, conversely, facilitate infection via immune suppression (e.g., Geldmacher & Koup 77 78 2012). The sequence in which pathogens infect an individual or 'priority effects' have been experimentally shown to be important in shaping co-infection dynamics in a variety of systems 79 80 (e.g., Hoverman et al. 2013; Halliday et al. 2017), yet are rarely demonstrated in non-81 experimental contexts. How priority effects and pathogen traits (e.g., transmission mode) affect 82 the nature and frequency of associations between endemic and epidemic pathogens, ultimately

shaping pathogen infra-communities is a knowledge gap that has significant consequences for
understanding patterns of infection (Munson *et al.* 2008; Telfer *et al.* 2010; Ezenwa & Jolles
2015; Halliday *et al.* 2017).

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Quantifying associations between pathogens from observational data and inferring interactions 87 from these patterns, however, is a methodological challenge (Fenton et al. 2014). Discriminating 88 between positive (i.e., two pathogens are more likely to occur together) or negative associations 89 (i.e., two pathogens are less likely to occur together) between pathogens in populations is 90 complicated by the short time window that a pathogen is shedding (and thus detectable with 91 92 molecular methods) and by potentially confounding host immune environments (Tompkins et al. 2011). This is particularly the case for microparasites where pathogen detection often relies on 93 serology, and, thus, without resampling the same individual, the precise timing of exposure 94 cannot be estimated. Detection of pathogens that form chronic infections may be more 95 straightforward as the infection is active for longer periods, but deducing pathogen associations 96 97 is difficult without extensive longitudinal data (Fenton et al. 2014; Hellard et al. 2015). 98 Identifying whether two pathogens are associated due to host-habitat preferences, the increasing likelihood of exposure with age, or are a product of a negative (e.g., competition) or positive 99 100 (e.g., facilitation) interactions is methodologically challenging (Poulin 2007; Johnson & Buller 2011; Fenton et al. 2014; Hellard et al. 2015; Clark et al. 2016). Identifying associations that 101 102 could represent candidate interactions based on observational data can not only provide a basis for experiments to test potential interactions but also provide novel insights into pathogen infra-103 104 community dynamics.

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106 Detecting associations between pathogens is also likely to depend on taxonomic and spatial 107 scales that are seldom considered (Araújo & Rozenfeld 2014; Stutz et al. 2018). Studies commonly aggregate pathogen data to genus level, but associations between pathogens can be 108 subtype or genotype-specific (e.g., Wejse et al. 2015; Benesh & Kalbe 2016; Brook et al. 2017). 109 110 For example, individuals infected with human immunodeficiency virus subtype 1 (HIV-1) are 111 four times more likely to become co-infected with tuberculosis compared to individuals with HIV-2 (Wejse et al. 2015). Beyond subtype or genus, genotype-specific associations have been 112 demonstrated in snails infected by trematodes (Louhi et al. 2015) and in rodents infected by 113

114 Bartonella bacteria (Brook et al. 2017). Infra-community dynamics are also likely to vary with spatiotemporal scale. In general, associations between free-living species are more apparent at 115 scales where interactions occur compared to broader spatiotemporal scales (Eltonian noise 116 hypothesis; Peterson et al. 2011; Araújo & Rozenfeld 2014), but it is unclear if this is true for 117 pathogens. Nonetheless, for cross-sectional datasets, important patterns may be missed unless 118 multiple spatio-temporal scales are considered (Ovaskainen et al. 2017). To overcome these 119 120 challenges, analytical approaches that can quantify associations between pathogens whilst controlling for potential confounding factors are required to assess the role of associations in 121 shaping pathogen infra-communities. 122

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Recent applications of network theory to parasite community ecology provide an opportunity to 124 125 move beyond the pairwise associations between two pathogens (Clark et al. 2016; Aivelo & Norberg 2018; Stutz et al. 2018). Network measures have frequently been used to study food 126 127 webs but are increasingly applied to pathogen infra-communities where nodes are pathogens, and edges represent pathogen co-occurrences within the host (Vaumourin et al. 2015). Networks are 128 129 modular if pathogens co-occur more frequently in particular groups, 'nested' if pathogens frequently share interaction partners across the network, or 'segregated' if the inverse is true 130 131 (Strona & Veech 2015; Ulrich et al. 2017). If, for example, networks are segregated, targeted 132 control of one 'keystone' pathogen may lead to co-extinction of other pathogens in a module 133 (Pedersen & Fenton 2007; Säterberg et al. 2013). If a network is nested, perturbations to the 134 pathogen infra-community may spread throughout the network (Griffiths et al. 2014).

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136 Although pathogen co-occurrence networks are valuable for quantifying broad structural 137 patterns, they do not account for environmental or host factors, pathogen traits or differences in 138 spatial or temporal scale. Joint species distribution models (JSDMs) fill this gap by simultaneously assessing environmental influences and interspecific co-occurrences across 139 multiple scales using hierarchical Bayesian mixed models (Warton et al. 2015; Ovaskainen et al. 140 141 2017). Here we use both co-occurrence networks and JSDMs to examine the structure of pathogen-pathogen networks and quantify pathogen associations while controlling for 142 environmental/host factors and scales. We include information on pathogen traits such as 143 144 transmission mode to assess what role they played in the distribution of each pathogen. We

collate ten years of cross-sectional data on endemic and epidemic pathogens in 105 African lions
(*Panthera leo*) as well as extensive host and environmental data from the Serengeti Lion Project
(SLP, Packer *et al.* 2005). The SLP datasets provide a unique opportunity to understand
pathogen co-occurrence networks in a wild population while controlling for group, individual
and environmental characteristics. We use this data to ask the following interlinked questions at
two levels of taxonomic resolution:

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- (I) To what degree is the pathogens' co-occurrence network of Serengeti lions nested orsegregated?
- (II) After accounting for environmental/host factors and spatio-temporal scale, is the type
  of endemic pathogen an individual is infected by early in life associated with
  exposure to epidemic pathogens later in life?
- 157 (III) Are there significant endemic-endemic or epidemic-epidemic pathogen co-158 occurrences?
- Because we could not directly determine the order of infection events from cross-sectional data 159 in isolation, we used age-prevalence relationships in combination with the natural history of each 160 161 pathogen to estimate probable timing of events. We describe an analytical pathway that can 162 assess broad network structure and quantify pathogen associations across multiple scales that can 163 generally be applied to understand infectious disease dynamics. The co-occurrence network 164 detects clusters of pathogen sharing amongst individuals and screens for disconnected nodes (pathogens that rarely co-occur with others), while the JSDM approach was used to quantify 165 pathogen-pathogen associations. To assess the plausibility of these putative interactions, we 166 167 compare our findings to similar mammalian pathogens in experimental studies. Detecting pathogen co-occurrences not only provides novel insights into pathogen infra-community 168 dynamics but also helps aide surveillance efforts in the field and generate testable hypotheses 169 that can be answered in laboratory experiments. 170

171 Methods

172 Pathogen data

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10 March 10

174 Serological testing and quantitative PCR (qPCR) were performed to detect endemic and epidemic pathogens from blood samples taken from lions in the Serengeti National Park, 175 176 Tanzania from 1984-1994. In total, 394 individuals were sampled throughout this period, but our 177 analysis was restricted to the 105 individuals tested for the full suite of ten pathogens (Table 1: pathogen natural history; Table S1: number of individuals tested per year included in the 178 analyses). Nomadic individuals (i.e. lions that were not resident in any pride) were excluded due 179 180 to the difficulty of assigning environmental variables (see Confounding variables below). Serological data on canine distemper virus (CDV), feline calicivirus, parvovirus, and coronavirus 181 has been published previously, except Rift Valley Fever (RVF) (Packer et al. 1999, see Table S2 182 183 for assay details). To detect RVF exposure we conducted a plaque reduction virus neutralizing test (PRNT) that quantified virus neutralizing antibodies from serum following Scott et al. 184

185 (1986) protocol.

186 We used qPCR to identify nucleotides for feline immunodeficiency virus (FIV<sub>Ple</sub>) and the protozoan pathogens in this study (Table 1). Three distinct subtypes of FIV<sub>Ple</sub> co-circulate in 187 188 Serengeti lions (Troyer et al. 2005, 2011; Antunes et al. 2008) and thus subtype specific qPCR was performed (see Troyer et al. 2004, 2005 for qPCR protocols). The resultant 300 base pair 189 190 sequences from the *pol* gene were aligned and assigned to 21 operational taxonomic 191 units/genotypes based on a 95% molecular similarity threshold (see Fountain-Jones et al. 2017 192 for details). Lions also commonly get infected by a rich protozoan fauna including Babesia and Hepatazoon genera. We developed quantitative PCR protocols using density gel gradient 193 194 electrophoresis to identify each protozoan species (see Munson et al. 2008).

195 We categorized each pathogen as likely endemic or epidemic in the lion population: endemic 196 pathogens were considered to be constantly circulating and often infecting the young while epidemic pathogens sweep through the population every few years infecting all age classes 197 (Packer et al. 1999; Penzhorn 2006; Troyer et al. 2011). Many of the pathogens have been 198 previously classified as endemic or epidemic (Packer et al. 1999). We supported our 199 200 classification with age-prevalence plots (Fig. S1) and we plotted yearly prevalence (Fig. S2) for 201 the pathogens not previously classified. Pathogens with a high prevalence at a young age ( $\leq 2$ y.o.) with little fluctuation across all years and age classes were considered to be likely endemic, 202 whereas an increasing age-prevalence relationship and high temporal variation were classified as 203

204 more likely to be epidemic in this population. Feline coronavirus can have epidemic and endemic cycles, and it is challenging to assess which form the individual was infected with from 205 206 serological data, but based on age-prevalence relationships we categorized coronavirus as an 207 endemic infection (Fig. S1). Further, we used patterns of age-prevalence to infer the potential timing of infections. As most individual lions were likely to be infected by the pathogens we 208 considered endemic within the first two years after birth (Troyer et al. 2011, Fig. S1), we assume 209 210 that endemic exposure typically occurred prior to exposure by an epidemic pathogen. We partitioned the endemic pathogen data into two sets based on taxonomic resolution (high and 211 medium). The high taxonomic resolution dataset encompassed FIV<sub>Ple</sub> genotype and *Babesia* 212 species data, whereas the medium resolution dataset aggregated FIV<sub>Ple</sub> subtype information and 213 Babesia data to genus level. 214

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#### Co-occurrence network

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217 We examined pathogen co-occurrence patterns to evaluate preferential associations among pathogens. We constructed co-occurrence networks for each taxonomic resolution as well as for 218 219 pathogens tested for using qPCR and by serology in cases combining both lines of diagnostic evidence led to altered network structure. To do so, we first built an  $m \times n$  matrix that described 220 presences/absences (i.e., occurrences) of both endemic and epidemic pathogens across individual 221 222 lions, where m was the number of individual lions and n the number of pathogens. By 223 multiplying it by its transpose, we then created a summary  $n \times n$  co-occurrence matrix that described, for each pair of pathogens, the number of observed co-occurrences across all 224 individual lions. Pathogens detected infrequently in this lion population were included in this 225 analysis to help screen for pathogens disconnected in the network. The co-occurrence matrix was 226 used to evaluate which pathogens were carried by the same individuals utilizing a modularity-227 based "greedy" approach (Clauset et al. 2004). Measures of modularity aim to determine the 228 adequacy of different classification schemes in representing clusters and divisions in datasets; 229 230 here, the clusters represented the co-occurrence of pathogens in individual hosts. Estimates of 231 modularity were calculated for each possible classification by comparing the expected fraction of pathogens co-occurrence to random co-occurrences (Newman 2006). The classification with the 232 highest modularity from all the generated classifications was selected. 233

We then computed a measure of network structure  $(\overline{N})$  and modularity index based on node overlap and segregation (Strona & Veech 2015; Ulrich *et al.* 2017).  $\overline{N}$  ranges from scaled from -1 (entirely segregated network) to 1 (entirely nested network). These analyses were performed in R using the 'igraph' and 'nos' libraries (Csárdi & Nepusz 2006; Strona & Veech 2015). The cooccurrence matrix was obtained from the incidence matrix using the graph.incidence and the bipartite.projection functions. The classification analysis was performed using the fastgreedy.community function in igraph (Csárdi & Nepusz 2006).

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# Joint species distribution modeling

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Joint species distribution models (JSDMs) are a flexible multivariate extension of generalized 243 linear mixed models that can examine how environment (and host) shape multiple species 244 simultaneously across biological scales (Ovaskainen et al. 2017; Björk et al. 2018). JSDMs can 245 quantify associations between species across scales using latent factor models to estimate 246 species-species covariance for each random effect (Ovaskainen et al. 2017; Björk et al. 2018). 247 We fitted JSDMs for both high and medium taxonomic resolution datasets, combining 248 information on environmental and host covariates as fixed effects (see Confounding variables 249 below for details), to the occurrence data for each of the pathogens. Pathogens detected fewer 250 than five times were excluded from this analysis leaving ten pathogens in the medium taxonomic 251 model and 17 in the high-resolution data set. Including pathogens with fewer than five 252 occurrences may lead to spurious associations (Ovaskainen et al. 2017). We fitted all the JSDMs 253 with Bayesian inference, using "Hierarchical Modelling of Species Communities" (Blanchet et 254 255 al. 2018). For each analysis, we modeled the response pathogen co-occurrence matrix using a 256 probit model based on the approach outlined in Ovaskainen et al. (2016). In contrast to the 257 network approach summary, the JSDM co-occurrence matrix is a product of the pathogen-topathogen variance-covariance matrix estimated for each random effect (e.g., pride-year) in the 258 259 model. Each random effect (and thus each estimated co-occurrence matrix) measures a 260 component of the variation in the response that is different than the other random effects and of 261 the set of explanatory variables (fixed effects) considered in the model. In our models, we added 262 individual (e.g., sex and age), pride, and environmental characteristics (see Confounding variables below) as fixed effects. Individual sampled, pride-year (i.e., which pride and year the 263

individual was sampled in), and year-landscape (i.e., what year was the individual sampled in the
Serengeti) sampled were added as random effects. As pathogen traits may shape the distribution
of each pathogen (e.g., similar environmental and host variables may shape tick-borne
pathogens), we included traits such as pathogen type (see Table 1) in each analysis. We utilized
the default priors (described in full detail in Ovaskainen *et al.* 2017) and ran the HMSC model
twice using 3 million MCMC samples (the first 300 000 of which being burn-in). Each run was
carried out using a different seed. Visual inspection of MCMC traces and the Gelman-Rubin

diagnostic calculated to assess convergence. In addition, we made sure that the effective sample
size (ESS) of each parameter was > 200.

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# 274 Confounding variables

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As part of the SLP, most of the individuals in this study have been regularly observed since birth (Mosser & Packer 2009). We selected 13 predictor variables that we thought were likely to be important for pathogen exposure and thus could confound possible associations patterns (Table 2). We included variables that captured individual variability (e.g., age at sampling), and pride characteristics including environmental variables (e.g., average vegetation cover of the pride's territory; see Table 2 for measurement details).

# 282 **Results**

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The Serengeti lions were exposed to an average of 5 pathogens (two epidemic and three endemic, SD = 1); one individual had been infected by 9 of 10 pathogens (based on medium resolution data, Fig. S3). Cubs between 1 and 2 y.o. were often already infected with an average of 4 pathogens (SD = 1), with one 1.5 y.o cub positive for 5. All lions were qPCR positive for at least one protozoan species, and 25% of them were infected by all four protozoans tested.

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# Pathogen co-occurrence networks are highly nested

292 The high taxonomic resolution summary network indicated a significantly nested architecture ( $\overline{N}$ = 0.74) with relatively low modularity (modularity index = 0.393, z = 3.307, p = <0.001) with 293 three clusters (Fig. 1a). The largest cluster (green nodes) included all of the protozoans, epidemic 294 pathogens, and some FIV<sub>Ple</sub> genotypes, whilst the remaining two clusters consisted of FIV<sub>Ple</sub> 295 genotypes (Fig. 1a). When we modeled networks based on diagnostic test, the general pattern did 296 not substantially change, with the exception that RVF clustered separately from the other viruses 297 detected using serology. In both network formulations, phylogenetically similar genotypes of 298 FIV did not cluster together (Fountain-Jones et al. 2017, see Fig. S5). In contrast, the medium 299 resolution network was completely nested with no modularity ( $\overline{N} = 1$ , modularity index = 0, z = 300 301  $\infty$ , p = 0) and no significant clusters (Fig. 1b).

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After accounting for environmental, individual and pride factors and scale, the JSDM analysis 305 306 identified strong associations between pathogens (Fig. 2) that were not detected in the summary co-occurrence network. Including individual, pride-year and landscape-year scales in our co-307 308 occurrence models was important as our ability to detect associations varied. At an individual and pride-year level, we detected strong associations between a small subset of epidemic and 309 310 endemic pathogens. FIV<sub>Ple</sub> B and H. felis were negatively associated with RVF (Fig. 2), and FIV<sub>Ple</sub> B was also negatively associated with parvovirus (Fig. 2a/S6a). However, these 311 associations could only be detected at medium taxonomic resolution. In contrast, at high 312 taxonomic resolution, we identified positive associations between *B. gibsoni* and RVF that were 313 314 not detected at medium-resolution.

Strong associations between endemic and epidemic pathogens

The strongest associations between endemic and epidemic pathogens were detected at the lowest spatial-temporal resolution (landscape-year). In the high taxonomic resolution model, pathogens separated into two groups with each group having a very similar association profile. One group was characterized by positive associations between the *Babesia* species, FIV<sub>Ple</sub> C2, CDV, and 319 parvovirus. The other group was characterized by positive associations between two FIV<sub>Ple</sub>

genotypes (C1 & B2), coronavirus, *H. felis* and calicivirus (Figs. 2c/S6c). There were strong

- negative associations between pathogens in each separate group (e.g., CDV and FIV<sub>Ple</sub> C1).
- 322 Generally, the same associations held in the medium taxonomic resolution models (Fig. 2c/S6c),
- but with exceptions. For example,  $FIV_{Ple}$  C1 and C2 had opposing association profiles, but as
- FIV  $_{Ple}$  C1 had a higher prevalence (Fig. S7), C1 had the same overall association profile as
- $325 \quad \text{FIV}_{\text{Ple}} \text{ C.}$

Associations between epidemic pathogens were rare. At the year-level, we detected positive 326 associations between CDV and parvovirus with both pathogens negatively associated with 327 calicivirus (Fig. 2c/S6c). In contrast, associations between the endemic pathogens were common, 328 329 but the nature of the associations also differed at each taxonomic scale. For example, in the medium resolution model, we detected a positive association between H. felis and FIV<sub>Ple</sub> C not 330 331 found in the high-resolution model indicating that FIV<sub>Ple</sub> subtype, but not genotype, was important for this association (Fig. 2b/S6b). Strikingly, we found that FIV<sub>Ple</sub> subtypes had 332 333 contrasting association profiles. At the individual level, FIV<sub>Ple</sub> B and C were negatively associated with each other, and FIV<sub>Ple</sub> C was positively associated with coronavirus, while 334 FIV<sub>Ple</sub> B was negatively associated with coronavirus (Fig. 2a/S6a). Both high and medium 335 taxonomic resolution JSDMs had reasonable explanatory power (Tiur  $R^2 = 0.381 \& 0.330$ . 336 respectively). In both models, the landscape and host factors that explained the distribution of 337 each pathogen were not predicted well by pathogen traits (See Fig. S8). See Fig. 3 for a summary 338 339 of all of the associations detected across scales from our cross-sectional data and Figs. S9/10 for model details. 340

# 341 Discussion

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Here we demonstrate non-random associations in the pathogens infecting wild African lions,
with both negative and positive associations detected between endemic and epidemic pathogens.
While there was minimal structure in the summary co-occurrence network (Fig. 1a), we
uncovered structure after accounting for scale and controlling for potentially confounding
environmental and host variables and scale (Fig. 2). Using age-prevalence relationships we could

assess the likely order of infection using cross-sectional data. We found that the particular 348 endemic pathogen an individual is infected by as a cub may have consequences for which novel 349 350 epidemic pathogen the individual is infected with later in life (Fig. 3). We emphasize that the approach used here can start to untangle pathogen infra-community relationships and identify 351 potential endemic-epidemic associations in wild populations. These can then be compared with 352 knowledge of pathogen pathogenesis and validated *in-vitro* in a laboratory setting. While clinical 353 354 or laboratory studies of co-infection in lions are rare for good reason, the associations we found 355 have clear precedence in similar pathogens co-infecting humans and represent plausible interactions. Our results not only provide new insights on pathogen community structure in the 356 357 Serengeti lions but also provide a valuable framework for exploring pathogen co-occurrence networks and infra-community dynamics. 358

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Co-occurrence networks were highly nested with relatively low modularity, particularly at a 360 361 medium taxonomic resolution. Nonetheless, RVF did cluster separately from the other pathogens tested via serology which potentially indicates that RVF, unlike the other epidemic viruses, has a 362 distinct epidemic cycle with most of the interacting partners being more chronic pathogens. This 363 is supported by the unique association profile detected in our JSDM analysis and is intuitive 364 given that RVF is the only mosquito-borne pathogen that we sampled. Even though we sampled 365 pathogens considered important for lion health, we lacked data on other potentially pathogenic 366 bacteria, helminths, and fungi that the lions were exposed to or potentially infected by. Further, 367 symbiont interactions can also be important in shaping pathogen dynamics (e.g., Halliday et al. 368 2017) and could be considered in pathogen infra-community studies. These additional taxa may 369 370 lead to further segregation in the network, as larger and more diverse networks typically show increased modularity and segregation (Thebault & Fontaine 2010; Sauve et al. 2014). Expanding 371 sampling to construct a more complete microbe and macroparasite network would also capture a 372 broader array of potentially facilitative and competitive associations (Ezenwa 2016; Aivelo & 373 Norberg 2018). 374

376 After accounting for environment, host and scale, we found that the endemic pathogens were 377 strongly associated with the epidemic pathogens and, based on mammalian lab-based 378 experiments, suggest that these patterns represent plausible interactions between pathogens. For example, we detected negative associations between endemic pathogens (FIV<sub>Ple</sub> B and H. felis) 379 380 and RVF after accounting for differences between individuals. Coinfections between bunyaviruses like RVF and retroviruses are likely common in humans and wildlife, though there 381 382 are surprisingly few studies addressing the topic. In contrast, relationships between dengue virus (a flavivirus) and HIV are relatively well understood. Flaviviruses and HIV share similar 383 immune receptors that can inhibit HIV replication and the molecular machinery used to do so 384 may be a viable way to control HIV infection (e.g., Xiang *et al.* 2009). Given the overall 385 structural similarity of flaviviruses and bunyaviruses (Hernandez et al. 2014), it is possible that a 386 similar mechanism underlies the association in lions between RVF and FIV<sub>Ple</sub> that we observed, 387 although we show that this association was subtype specific. If this was true, RVF might inhibit 388 FIV<sub>Ple</sub> B infection— counter to our assumption that endemic pathogens in our system infected 389 each individual first (Fig. 3). 390

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The greatest number of associations between epidemic and endemic pathogens were detected 392 when we included differences across years (year-landscape scale) in our analysis. These 393 associations could represent plausible facilitative or competitive interactions. CDV and Babesia 394 are well known to interact with high levels of Babesia infection magnifying the impacts of 395 396 consequent T cell depletion caused by CDV infection leading to mortality of nearly 40% of the lion population in 1994 (Munson et al. 2008). We found that all tick-borne hemoparasites 397 showed positive associations with CDV including B. leo (with insert) despite its low prevalence 398 in 1993/4 (Fig. S9). Parvovirus was also positively associated with CDV, but this was likely due 399 to similarities in timings of epidemics with a parvovirus epidemic in 1992 just before the 1994 400 CDV epidemic (Packer et al. 1999). Parvoviruses are also immune suppressive, and so the timing 401 402 of the parvovirus outbreak may also have contributed to the CDV/Babesia-induced mortality. The general negative relationship between FIV<sub>Ple</sub> C and CDV/Babesia supports the theory that 403 404 individuals infected by subtype C were more likely to die in the consequent *Babesia*/CDV

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Our approach detected strong associations between the endemic pathogens also. For example, 408 there were opposing associations between the FIV<sub>Ple</sub> subtypes and coronavirus (Fig. 2). Negative 409 410 associations between retroviruses and coronaviruses are rarely reported, yet there are plausible molecular pathways. HIV-1 and human coronaviruses (HCoV) share remarkably similar binding 411 receptors (Chan et al. 2006) and some mild HCoV strains are even considered a viable vaccine 412 against HIV (Eriksson et al. 2006). This may explain the negative association we detected for 413 414 FIV<sub>Ple</sub> B and coronavirus but does not explain the positive association between FIV<sub>Ple</sub> C and coronavirus we detected across scales. The mechanism driving FIV<sub>Ple</sub> subtype specific 415 relationships with coronaviruses are unclear, and as coronaviruses infecting lions are also likely 416 to be genetically diverse, examining the genetic structure of coronavirus may help untangle these 417 418 associations further. In contrast, competitive associations between HIV strains are well 419 characterized with HIV-1 found to outcompete HIV-2 for blood resources (Ariën et al. 2005). For FIV<sub>Ple</sub>, even though co-infection is relatively common (Troyer *et al.* 2011) competition 420 between subtypes could be important as there is anecdotal cell culture evidence that FIV<sub>Ple</sub> B can 421 propagate more rapidly than FIV<sub>Ple</sub> C (Melody Roelke, unpublished data). 422

outbreak (Troyer et al. 2011). Thus, this negative association may not be due to competition

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There were also contrasting associations between the protozoan species. For example, the 424 distribution of *B. felis* was not shaped by any other protozoan and in general, had a narrow 425 association profile (Fig. 2), unlike the other *Babesia* species. For the individuals co-infected by 426 protozoans, associations involving B. felis were also common, whereas co-infections involving 427 H. felis and the other Babesia species varied in prevalence and composition (Fig. S9). Even 428 429 though *B. gibsoni* and *B. felis* show similar age prevalence profiles (Fig. S1), the prevalence of B. felis over time was relatively stable compared to the other protozoa (Fig. S10). Differences in 430 the host range for individual *Babesia* species and potential host differences in virulence may 431 partially explain these patterns. For example, B. felis has only ever been detected in felids, 432 whereas B. gibsoni has a much broader host range including canids (Penzhorn 2006). Generalist 433

between pathogens but rather to mortality.

pathogens may have greater pathogenicity as there can be reduced selective restraint on virulence
particularly in 'dead end' hosts (Woolhouse *et al.* 2001). If more pathogenic species are more
likely to interact with other pathogens compared to less virulent pathogens is an open question in
disease ecology. Importantly, patterns like these would be missed without incorporating highresolution pathogen data.

There are, however, limitations to this approach. The inability to distinguish mortality or 439 440 correlated exposure (i.e., an individual is infected by multiple pathogens in the same transmission event) from negative associations is one of them, and careful interpretation of 441 negative associations is necessary. Incorporating approaches such as structural equation models 442 that explicitly include potential mechanisms that underlie candidate pathogen associations 443 444 (Carver et al. 2015) could be a valuable additional step in future pathogen network studies. Another weakness is the inability to estimate the timing of these infections more precisely. For 445 446 example, the negative association between RVF and H. felis could be due to temporal differences when ticks and mosquitoes emerge after rains. Years with higher rainfall increase mosquito 447 448 abundance thus increasing RVF prevalence (Fig. S2) whereas ticks emerge en masse when rains 449 follow a dry period potentially increasing H. felis prevalence (Munson et al. 2008). As rainfall 450 was calibrated to the year of sampling rather than the age of infection (which could differ) the 451 JSDM approach could not capture this variation. Studies using longitudinal data to quantify 452 associations using a similar framework to ours (e.g., Telfer et al. 2010; Henrichs et al. 2016) will be beneficial as they are likely to provide more robust estimates of the order of infection in wild 453 454 populations. Furthermore, we cannot quantify the importance of these associations in shaping pathogen distribution across scales compared to processes such as host density. Lastly, 455 456 incorporating immune function and host resources in both the summary network and JSDM 457 analyses are likely to provide mechanistic insight into pathogen network structure (Griffiths et al. 2014). Higher resolution pathogen traits, such as duration of infection, are likely to provide 458 459 further mechanistic insight into how and why pathogens co-occur as they do in free-living communities (Ulrich et al. 2017). However, given the daunting complexity of pathogen infra-460 community dynamics, our two-step approach can assess broad network structure and identify 461 462 useful candidate interactions between pathogens thereby reducing some of this complexity.

The high frequency of co-occurrence and co-infection in lions – and the potential for specific 464 associations to cause population decline – highlights the importance of understanding pathogen 465 466 associations. The summary lion pathogen co-occurrence network was highly connected with both positive and negative associations between endemic and epidemic pathogens. Our findings 467 indicate that the lion pathogen infra-community is influenced by a number of ecological factors 468 469 and associations between pathogens. We identify useful associations between pathogens thereby reducing some of this complexity. More broadly, our work demonstrates how different network 470 approaches can be combined to gain insights into the ecological factors underlying pathogen 471 associations and how this can be applied to the study of pathogen communities in wildlife 472 populations. In addition to these biological insights, the study highlights several critical areas for 473 methodological improvement that can currently limit robust inference of pathogen associations 474 475 from cross-sectional serological and qPCR data. Addressing these limitations is timely, given the ongoing threat of wildlife population decline, creating a need to integrate better molecular, 476 477 ecological and network information for disease control.

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662	<b>Table 1:</b> Traits of both endemic and epidemic pathogens in this study.

Pathogen	Туре	Trans.	One	Immune	Exposure	Test
		mode	host?	sup.	timing? <sup>*</sup>	type
Data type	(binary)	(categorical)	(binary)	(binary)	(binary)	(binary)
EPIDEMIC						
Feline calicivirus	Virus	Direct/env	Ν	NE	Epidemic	Serology
(calicivirus) †					year	
Canine distemper	Virus	Direct	Ν	Yes	Epidemic	Serology
virus (CDV)					year	
Feline	Virus	Vertical,	Ν	Yes	Epidemic	Serology
panleukopenia		direct/env			year	
(parvovirus)						
Rift valley fever	Virus	Vector	Ν	Yes	Throughout	Serology
(RVF)		(mosquito)			life#	
ENDEMIC						

Feline enteric	Virus	Vector	Ν	U	Epidemic	Serology		
coronavirus					year			
(coronavirus) †								
B. gibsoni	Protozoa	Vector (tick)	Ν	NE	< 2 y.o	qPCR		
B. leo with	Protozoa	Vector(tick)	Ν	NE	Throughout	qPCR		
insertion					life			
B. felis	Protozoa	Vector (tick)	Ν	NE	< 2 y.o.	qPCR		
Hepatozoon felis	Protozoa	Vector (tick)	Ν	NE	< 2 y.o.	qPCR		
Feline	Virus	Vertical/	Y	Yes	< 2 y.o.	qPCR		
immunodeficiency		direct						
virus								
FIV <sub>Ple</sub> A, B, and C								
and FIV								
genotypes A1, B1-								
12, C1-C8								
Trans. mode : Transmission mode (all pathogens can be horizontally transmitted). Immune sup.: Pathogen can								

664 suppress the immune system. Vertical: Vertical transmission is also possible. Env: Environmentally persistent.

665 Direct: Transmission through host contact. Immune sup.: Immune suppression.\*: Likely time of exposure †:

666 Determined by age-prevalence relationships (see *Methods* and Fig. S1) but can have endemic or epidemic variants.

667 U: Unknown NE: No evidence. #: More likely after heavy rainfall (Fig. S2).

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**Table 2:** Details of the individual, pride-level and environmental predictors used in the joint

- 670 species distribution models to help account for potential confounding factors. All variables were
- 671 calculated based on the year of sampling.

Predictor	Туре	Measurement details	Data
Sex	Individual	Male or female.	SLP data
Age	Individual	Age of lion when sampled (days).	SLP data
Number of	Individual	Number of prides an individual has	SLP data
immigrations		immigrated into prior to sampling.	
Pride or	Individual	Was the male involved in a coalition	SLP data
coalition male?		occupying multiple prides (binary)?	

Group size	Pride	Average number of individuals in	SLP data
		pride two years <sup>†</sup> prior to sample	
		collection.	
Despotic	Pride	Was the pride considered despotic at	SLP data
		time of sample collection?	
Territory size	Pride	Based on location data over a two-year	SLP data
		period based on utilization-	
	_	distribution curves with a 75% kernel.	
Territory	Pride	What percentage of territory size	SLP data
overlap		overlapped with other prides.	
Habitat quality	Pride	Pride habitat quality score calculated	(Mosser et al. 2009)
U		across a two-year period.	
Number of	Pride	Number of individuals in neighboring	SLP data
neighbors		prides. Neighboring prides had	
		territory overlap.	
Yearly rainfall	Environmental	Yearly rainfall experienced in each	(Sinclair et al. 2013)
Ω		pride territory based on weather	
		stations in the plains and woodlands.	
Average	Environmental	Average vegetation cover across the	(Reed et al. 2009)
vegetation cover	-	prides territory based on a 75% kernal.	
Soil pH	Environmental	Average pH throughout the pride's	World Harmonized Soil
	-	territory based on a 75% kernel.	Database (FAO &
C			IIASA 2009).

\* We calculated this predictor two years prior to sampling to account for differences in individual status at a

673 potential time of exposure or infection (e.g., individuals that had just immigrated into a pride when sampled were

- 674 considered nomads as exposure or infection was likely to have occurred previously). †: We averaged over past two
- 675 years to reduce the variability in pride counts as exposure was unlikely to have happened during the sampling year.

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Fig. 2: Pathogen-pathogen associations detected at (a) individual, (b) pride-year and (c) 686 landscape-year level after controlling for individual, pride, and environmental variables in high 687 and medium taxonomic resolution models. Blue represents negative correlations and red 688 indicates positive associations. Only associations with posterior coefficient estimates  $\geq 0.4$  with 689 95% credible intervals that do not cross 0 are shown. The light red line indicates the assocation 690 between *H*. *felis* and CDV that was  $\ge 0.4$  in the medium resolution model but was below the 691 threshold (0.38) in the high resolution model. Pathogens in bold and in boxes are the epidemic 692 viruses (all other pathogens are likely endemic). This figure was drawn using the R package 693 'circleplot' (Westgate 2016). See Fig. S6 for association matrices and Figs. S9/S10 for covariate 694 695 partitioning and effect size.



Fig. 3: Summary of the strong positive (red line/arrows) and negative (blue lines/arrows)
associations between endemic (grey circles) and epidemic (orange circles) pathogens in the
Serengeti lions; dark-grey borders indicate protozoa. The direction of the red or blue arrows
indicates the potential sequence of infection events. The black arrow along the X-axis represents
age; the circles reflect the ages when lions were likely to be infected by each pathogen (based on

- age-exposure data rather than longitudinal data, see Fig. S1). Dashed circles indicate major co-
- 704 occurrence clusters identified at the landscape-year scale.

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