
Stock assessment on fishery- dependent data: Effect of data quality and parametrisation for a red snapper fishery

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Abstract :

Data availability, and unreported and unregulated fishing are significant obstacles to evaluating stock status, especially in tropical areas. Limitations in data quantity and quality can lead to model misspecification and erroneous data treatments, potentially causing important changes in model outputs and subsequent management implications. Red snapper *Lutjanus purpureus* (Poey) in French Guiana provides an example of a stock with a long-time series of fishery-dependent data subject to large uncertainty. A flexible catch-at-age model (Stock Synthesis) was applied to the available data and compared to an historically applied assessment approach. Inter-model variability based on different model specifications and data treatments were compared to identify better the status of the resource. Results showed that a major source of uncertainty in the model was the inclusion of a catch-per-unit-effort abundance index with questionable ability to track abundance. The Stock Synthesis model provided a more flexible and viable method than the virtual population analysis approach. Despite large uncertainty, models depicted a similar trend with a notable stock depletion in the late 1990s but with two distinct biomass trends in more recent years depending on the treatment. To reduce uncertainty and preserve this important economic resource, new data collection programmes and management policies are needed.

Keywords : data-limited, French Guiana, *Lutjanus purpureus*, stock synthesis

25 **Introduction**

26 Quantitative fishery stock assessments look to produce data-driven estimates of population
27 abundance and dynamics to inform management decisions (Hilborn & Walters, 2013).
28 Measurement error associated with a variety of data types and parameters, and natural
29 process variability in population dynamics all lead to uncertainty in analytical outputs (Francis
30 & Shotton, 2011). This uncertainty can hinge greatly on the quality and availability of the data
31 (Chen, Chen, & Stergiou, 2003). Despite the potential for data-derived biases in assessments
32 that could lead to management failures, improvements in quality and data availability are often
33 limited by resources like money, time and available expertise (Chen et al., 2003). Developing
34 harvest strategies based on limited data without waiting for extensive data sets that may never
35 materialize is critical to responsive and responsible natural resource management (Dowling et
36 al., 2015).

37 Data limitation is a global problem, particularly in tropical regions (Amorim, Sousa, Jardim, &
38 Menezes, 2019). In addition to the challenges of monitoring legal activities, unreported and
39 unregulated fishing presents an additional significant challenge to informing stock
40 assessments (Cawthorn & Mariani, 2017). For example, unreported catches and effort can
41 lead to severely biased estimates of biomass and other model outputs (Omori, Hoenig,
42 Luehring, & Baier-Lockhart, 2016). This creates a mixed situation of partial coverage in data
43 streams resulting in uncertainty that deserves respect and acknowledgement, but such
44 uncertainty is an insufficient reason to avoid using science to inform management (Dowling et
45 al., 2016). Unfortunately, numerous valuable fishery resources, especially in tropical areas,
46 remain unevaluated.

47 The red snapper (*Lutjanus purpureus*; Poey, 1866) fishery in French Guiana constitutes a good
48 example of a data-limited fishery where, despite the availability of long time series of fishery-
49 dependent data, information on the ecology of the species is still poorly understood and data
50 gaps remain. The commercial handline red snapper fishery in French Guiana is predominantly
51 performed by Venezuelan boats under a licensing system introduced in the 1980s by the

52 French government and now under EU authority. The licence agreement requires boats to sell
53 75% of their catch to a processing factory in French Guiana, while the other 25% can be sold
54 abroad. Controls at sea and at landing sites exist, however no information is available on the
55 catch sold beyond French Guiana borders. Moreover, this fishing activity is mostly focused on
56 smaller fish since the international market demands plate-sized fish typically below the size of
57 maturity. This type of size-selective fishery can lead to age truncation if fishing mortality is high
58 (Brunel & Piet, 2013; Reddy et al., 2013). Even if the catch of large fish declines, high fishing
59 intensity on small individuals can threaten population sustainability (Reddy et al., 2013).

60 *L. purpureus* is particularly vulnerable to fishing pressure due to its behaviour and general life
61 history characteristics (slow-growth, late maturity and seasonal spawning aggregations;
62 (Manickchand-Heileman & Phillip, 1996). Red snapper is known to aggregate for spawning, a
63 behaviour that can lead to hyperstable signals of population density and overestimation of the
64 stock size if not accounted for in stock assessments (Erisman, Apel, MacCall, Román, & Fujita,
65 2014). Additionally, the specific life history of *L. purpureus* is not well understood, creating
66 significant uncertainty in the use of biological parameters (e.g., growth, natural mortality, and
67 reproduction) in French Guianan waters (Rivot, Charuau, Rose, & Achoun, 2000).

68 Red snapper in French Guiana provides an example of a stock with multiple data sources (e.g.,
69 catch, fishery-dependent index and biological compositions) of limited quality and uncertain
70 life history values. Consequently, it is critical to compare several possible model specifications
71 and data treatments in order to account for uncertainty in the estimation of management
72 quantities in any stock assessment. This work takes up the challenge of assessing *L.*
73 *purpureus* by: 1) investigating stock assessment uncertainty based on limitations in the fishery-
74 dependent data and life history inputs using a flexible statistical catch at age approach (i.e.,
75 the Stock Synthesis (SS) modelling framework) and 2) comparing results from the SS model
76 to that of a Virtual Population Analysis (VPA) approach that has historically been used to
77 assess the stock. By accounting for the uncertainty in sources of data and inputs, the major
78 sources of model output uncertainty are identified and quantified for management

79 consideration, while using the flexible SS framework may provide a more advantageous
80 modelling environment compared to the more rigid VPA approach.

81 **Material and methods**

82 **Fishery catch data**

83 In French Guiana, red snapper is mostly fished by Venezuelan hand-liners between 30 and
84 200 m depth. The hand-line fishery is estimated to have started around 1960, with some
85 information on landings beginning in 1976, with the most reliably data recorded from 1985
86 onward when Ifremer (Institut Français de Recherche pour l'Exploitation de la Mer) started a
87 fisheries information database (Tous, 1988). Before 1988, other fishing activities such as
88 trawling were also targeting red snapper, but catches were not monitored (Prevost, 1989; Tous,
89 1988).

90 In 1984 a licencing system was implemented requiring Venezuelan boats to sell a fixed
91 percentage of their catch in French Guiana (50% for 1984 and 75% from 1985 to present). In
92 addition to the main hand-line fishery, a few boats coming from the French Antilles islands
93 occasionally fished French Guiana waters with fish traps (e.g., in 2019, less than 70 tons or
94 2.6% of total yearly catches). Additionally, bycatch in shrimp trawlers takes small (between 8
95 and 30cm) red snapper (i.e., in 2007 about 100 tons, or 6% of total catches for 23 trawlers),
96 but little information on the historical time series of these catches is available (Caro & Lampert,
97 2011). Currently, only 10 shrimp trawlers remain, likely reducing the amount of red snapper
98 bycatch. Considering the high uncertainty of the landings data, especially at the beginning of
99 the time-series, and the need to correct for these missing catches, landings data were
100 expanded by 25% for all years to estimate total removals from the red snapper population (Fig.
101 1). No information was available to hypothesize any temporal changes in the expansion value.

102 **Abundance data**

103 Fishery-dependent catch per unit effort (CPUE) indices were available from 1986 to 2018 (Fig.
104 1). CPUE were calculated by dividing the total annual catches by the total annual number of
105 days at sea estimated from logbooks and/or vessel monitoring systems (VMS) data (tons of

106 catch*days at sea⁻¹). CPUE were not standardized since no historical information in changes
107 of the fishing techniques or other factors were available. This lack of standardization adds
108 uncertainty in the application of this index, but it is the only index available.

109 **Biological data**

110 Length composition (fork length in cm) data were available from 1986 to 2019. Length is
111 routinely measured by observers at landing sites in Cayenne according to the framework of
112 the fisheries information system (SIH) implemented by Ifremer. The available data set is
113 obtained from a monthly sampling plan that subsamples boats landing red snapper. The
114 sample size has changed over the years following changes in the fishery and improvement in
115 the statistical analysis to try and optimize the sample size. The length frequency of the
116 subsample was therefore expanded to match the 25% expansion in landings (to account for
117 the animals fished in French Guiana waters but landed abroad).

118 **Life history relationships and values**

119 Natural mortality was assumed constant across ages and time. Individual growth is modelled
120 as a Von Bertalanffy function, fecundity was modelled as proportional to weight, and a
121 Beverton-Holt stock-recruit relationship was assumed. Life history values were fixed in the
122 reference model and were obtained from literature sources (Table 1). Exploration of
123 uncertainty in natural mortality (M) and the Beverton-Holt steepness parameters (h ;
124 recruitment compensation, or average recruitment of a population reduced to 20% of unfished
125 levels relative to average recruitment of the unfished population) are described in the next
126 section on sensitivity analysis.

127 **Model description and specification**

128 The assessment was conducted using the Stock Synthesis (SS version 3.30.13.02) framework
129 that uses maximum likelihood estimation (MLE) to obtain values and calculate asymptotic
130 uncertainty for estimated parameters and model outputs (Methot Jr & Wetzel, 2013). The
131 model is configured as one sex as females and males were assumed to have the same life

132 history parameters. Fishery-dependent data (catch, CPUE, and length compositions) were
133 specified as one fleet with dome-shape selectivity as the largest individuals are not taken in
134 the fishery, a parameterization choice confirmed by fishermen. A selectivity time block was
135 applied with a break implemented after 1996 and 2018 to better account for a possible change
136 in fishing practice (targeting smaller individuals to adapt to market demand) as suggested by
137 local fishermen that changed the size composition of the landed fish. The model with a time
138 block in selectivity improved model fit to the length compositions (see Appendix B and C).
139 Catch in metric tons was assumed known while the CPUE index assumed a lognormal error
140 with a standard deviation of 0.3 for all years. Length composition data were modelled with 2
141 cm length bins between 15 and 85 cm, and relative sample sizes among years were
142 determined by the samples by trip weighted by catch. The list of the parameters used in the
143 reference model is provided in Appendix A. The data and model outputs were summarized
144 using the r4SS package (<https://github.com/r4ss/r4ss>). Additional data weighting for lengths
145 and CPUE were unnecessary given the model fit (see Appendix C).

146 **Sensitivity analysis and likelihood profiles**

147 Model sensitivity to parameter uncertainty was explored via likelihood profiles—the fixing of
148 the model to various values of a specific parameter to see how model fit and derived outputs
149 change. Likelihood profiles demonstrate the amount of information (measured by the changing
150 likelihood metric) contained in the data for the featured parameter. Using the negative log
151 likelihood metric, any value outside of 1.96 units from the maximum likelihood estimate (MLE)
152 is considered significantly less supported by the data. The spread of model outputs within the
153 interval of significant data support therefore provides a measure of uncertainty in model output
154 based on parameter input. To demonstrate how model output changes across profiled
155 parameter values, three model outputs were considered: 1) initial spawning output (SO_0), 2)
156 terminal year spawning output (SO_{2018}), 3) the stock status in the terminal year (SO_{2018}/SO_0).
157 Comparing the information content of a particular parameter value to the associated model

158 output allows a mapping of model information (i.e., data) to sensitivity in the model output (i.e.,
159 results). Likelihood profiles were conducted for the following two parameters:

160 Natural mortality

161 Natural mortality is one of the most influential and difficult parameters to estimate in fisheries
162 stock assessment (Lee, Maunder, Piner, & Methot, 2011). Stock assessments often use an
163 external estimate of M as a fixed value, but may also estimate M within the model. Estimating
164 M depends on other model specifications (e.g., having at least one fishery with asymptotic
165 selectivity) and necessitates an exploration of model performance (Brodziak, Ianelli, Lorenzen,
166 & Methot Jr, 2011; Lee et al., 2011).

167 The range of M values used in the likelihood profile were defined by first estimating M indirectly
168 using meta-analytical and empirical methods based on life history parameters. “The Natural
169 Mortality Tool” (http://barefootecologist.com.au/shiny_m) application was used to access to
170 many different empirical M estimators. The methods based on maximum age (Hamel et al.,
171 2015; Then, Hoenig, Hall, & Hewitt, 2015) and on the von Bertalanffy K parameter (Alverson
172 & Carney, 1975; Jensen, 1996, 1997; Zhang & Megrey, 2006) and FishLife (Thorson, Munch,
173 Cope, & Gao, 2017) estimates were selected. M estimates varied between 0.09 and 0.46 year⁻¹
174 with a median value of 0.39 year⁻¹. These values were also compared to that of Rivot et al.
175 (2000) who compared three different estimation methods for French Guiana red snappers
176 suggesting that M ranged from 0.18 to 0.61 year⁻¹ (Pauly & Moreau, 1997; Ralston & Polovina,
177 1987; Rikhter & Efanov, 1976), with 0.29 year⁻¹ considered the most plausible. A likelihood
178 profile range of M from 0.10 to 0.60 year⁻¹ at an interval of 0.05 was defined using both of the
179 above sources (The Natural Mortality Tool and Rivot et al. 2000).

180 Steepness

181 It is common in stock assessments to define the functional relationship between spawners and
182 recruits using the reparameterized Beverton-Holt function (Mace & Doonan, 1988) where
183 steepness (h) is a key parameter. Steepness technically ranges from 0.2 to 1 in the Beverton-
184 Holt model, though values below 0.3 are often deemed unsustainable (He, Mangel, & MacCall,

185 2006). A higher h value loosens the relationship between stock and recruits, producing higher
186 productivity at smaller stock sizes. A value of $h=1$ essentially decouples the stock-recruit
187 relationship (Mangel et al., 2013; Shertzer & Conn, 2012). Steepness defines some
188 management quantities (e.g., MSY and F_{MSY}), but direct estimation requires contrast in the
189 data at low and high population sizes.

190 Externally-derived steepness values are much more commonly used, and come from life
191 history parameters and meta-analyses on ecologically similar species (Shertzer & Conn,
192 2012). The R package “FishLife” (<https://github.com/James-Thorson-NOAA/FishLife>; Thorson
193 (2020)) was used to specify h (0.7) for *L. purpureus*, and the subsequent likelihood profile
194 range of h was 0.40 to 1 with an interval of 0.05.

195 **Uncertainty in length data**

196 We explore the impact of bias in unsampled lengths from the unreported international fishery.
197 Unfortunately, no data is available on the proportion of fish sold abroad, but local fisherman
198 indicate bigger fish are typically landed for markets outside French Guiana. To test this
199 hypothesis, length compositions of the non-monitored landings (assumed to represent 25% of
200 the total catch) from 1991 to present were modified to include individuals larger than 40 cm
201 following the average length distribution composition for years 1986-1991 (period when larger
202 individuals were fished). This model was then compared to the reference model using only
203 sampled lengths.

204 **Uncertainty in CPUE**

205 The available raw CPUE data used in this study were exclusively derived from fishery-
206 dependent time series and were non standardized since little information are available on
207 sampling conditions, fish biology and movement patterns, or on changes in fishing behaviour.
208 To better understand the influence of the CPUE index on model outputs, the reference SS
209 model was compared to a model with no CPUE index, thus relying only on catches and lengths
210 as inputs. The assumption of linearity between CPUE and abundance was investigated by

211 estimating the exponent of a power function relationship between the CPUE index and the
212 catchability (Hilborn & Walters, 2013; Methot, 2009).

213 **Comparison to VPA**

214 The *L. purpureus* stock was first assessed in 2012 by applying a VPA on commercial length
215 frequency data from 1986 following a von Bertalanffy growth relationship and assuming a
216 maximum age of 13 years (Lampert, 2012). VPA uses a backward projection to estimate
217 recruits with no stock recruit relationship, while SS assumes the Beverton-Holt stock-
218 recruitment relationship. The SS model also assumes length variability at age, whereas the
219 length-age relationship in the VPA was taken straight from the von Bertalanffy curve. The VPA
220 model was constructed following the equations in example 18 of Sparre and Venema (1998).
221 A plus-group was employed for the last age group. F (and Z) are age-specific with the plus
222 group applying a constant average F value. The VPA model does not explicitly specify
223 selectivity. The VPA model did not consider the CPUE data and applied a constant fishing
224 mortality by cohorts (averaged over the most recent 5 years) to estimate stock biomass.
225 Natural mortality in the VPA model was fixed at 0.29. The VPA model was run again with the
226 most recent data and main outputs (total biomass, recruitment, spawning biomass and relative
227 spawning biomass relative to the first year of the model) were compared to the SS model.

228 **RESULTS**

229 **Reference Model**

230 The reference SS model of the red snapper shows a stock in initial decline, but in recent
231 years increases in biomass despite increasing catches (Fig. 2). Recruitments are at its
232 highest post-2000, when the CPUE time series shows a steady increase. Current relative
233 stock status is very high and well above what would be considered maximum sustainable
234 biomass (Fig. 2).

235 **Likelihood Profiles**

236 Natural mortality

237 The model tends to support higher values of natural mortality (Fig. 3, likelihood panel), but the
238 amount of information in the model on natural mortality is very limited. Most of the information
239 comes from the assumed prior on M when looking at the likelihood components in the profile
240 (Appendix C). Initial and final spawning output are very sensitive to the assumption of lower M
241 values (Fig. 3). Despite the sensitivity in the absolute biomass measures, the relative biomass
242 was similar across the full profile (Fig. 3). This illustrates a situation where the model is poorly
243 informed on the absolute biomass of the stock, but the current stock status is robust to changes
244 in perception of M and indicative of a high stock status.

245 The SS-estimated M value was particularly high (0.46 year^{-1}) and probably unrealistic for *L.*
246 *purpureus* given the life history and lack of information on M contained in the data (Fig. 3). For
247 this reason, the median value of 0.39 year^{-1} from the nine empirical estimation methods was
248 fixed and assumed for both sexes in the reference assessment model.

249 Steepness

250 The steepness likelihood profile showed the available data had no information on the
251 steepness value (Fig. 4). Biomass changed non-linearly to steepness, with higher biomass at
252 lower steepness values, a typical result when looking across steepness values. Relative stock
253 status, while somewhat sensitive to the value of h , was consistently high across all steepness
254 values given the other data and parameter specifications in the reference model.

255 **Uncertainty in length data**

256 The inclusion of larger individuals on the length compositions resulted in slightly different
257 estimated selectivity parameters in the two-time periods (Appendix E) that result in large
258 overlap in biomass estimates between the length composition treatments (Fig. 6). The small
259 differences between models are highlighted by slightly larger biomass estimates, higher
260 relative stock sizes and lower fishing pressure in the model including larger individuals, though
261 well within the bounds of uncertainty of the reference model.

262 **Uncertainty in CPUE**

263 Removing the CPUE index strongly affected model output, resulting in a more pessimistic
264 situation for both ending biomass and subsequent stock status (Fig. 7 and Appendix D). The
265 scale of the initial population biomass was not sensitive to inclusion of the CPUE index, but
266 the final biomass is sensitive, pointing to the importance of the CPUE index as a source of
267 current stock status information. Whether this data sets contains an unbiased signal relative
268 to noise regarding the trend in the population is a critical assumption when interpreting these
269 results.

270 The SS model using a power relationship between the CPUE index and the catchability
271 suggests hyperdepletion in the raw CPUE (estimated catchability power value of 2.62). The
272 model outputs also incorporated more uncertainty relative to the reference model but the trends
273 were similar (Appendix F). Any interpretation using the raw CPUE index should be considered
274 with enormous caution.

275 **Comparison to VPA**

276 SS outputs for the reference model and the model without CPUE were compared to the results
277 from the VPA model (Fig. 2). As previously demonstrated, these two specifications of the SS
278 models differ mostly in years after 2001. Before 2001, the VPA model showed relatively lower
279 biomass levels compared to the SS outputs, though both models suggest the lowest biomass
280 was in the early 2000s (Fig. 2). From about 2010, the outputs of SS model without CPUE
281 (recruitments, SSB and total biomass) resulted closer to the VPA estimation. On the other
282 hand, the relative stock status as defined by the first year of the time series ($SSB_{current}/SSB_{1986}$)
283 indicates a larger decline for SS model without CPUE compared to the VPA. Interestingly, the
284 VPA uses only the length data, not the CPUE, yet still shows recovery, whereas the SS with
285 no CPUE scenario shows a persistent decline.

286 Given the historical VPA assumes a lower M value compared to the SS models, an additional
287 VPA model with the same M value used in the SS model was performed. This sensitivity did

288 not result in enough change in the VPA model to account for the different biomass scales
289 between the VPA and SS models.

290 **Discussion**

291 When applying complex stock assessment models in data-limited situations, it is important to
292 have the flexibility to explore major axes of uncertainty and alternative model specifications,
293 not rely on the output of just one model. SS is a powerful and flexible modelling framework
294 accommodating many ways of exploring uncertainty, including data inputs and major life
295 history parameter exploration. Sensitivity analysis can be performed on several assumptions
296 (e.g. growth parameters or selectivity shape), but results can be difficult to interpret if the
297 probabilistic statements for the different values are unknown (Maunder & Piner, 2015). Natural
298 mortality and particularly steepness can be difficult to estimate in stock assessments, as both
299 benefit from contrast in the data. Though empirical estimators for M are available, their
300 imprecision (Kenchington, 2014) requires further characterization of uncertainty outside one
301 model specification (i.e., using only one value of M). Sensitivity analyses are recommended to
302 test for the robustness of model outputs to parameter and data choices and offer a fuller
303 representation of uncertainty and effects of model misspecifications (Brooks & Deroba, 2015).
304 Erroneous estimation of M can lead to over- or underestimates of stock biomass and status,
305 poorly informing management of the resource (Kenchington, 2014). Mortality rates for *Lutjanus*
306 species reported in literature from Florida to Brazil range widely from 0.11 to 0.49 year⁻¹
307 (Arreguín-Sánchez, Munro, Balgos, & Pauly, 1996; Burton, 2002; Rivot et al., 2000; Topping
308 & Szedlmayer, 2013). Our likelihood profile and sensitivity analysis showed the largest
309 changes in model output with low M values. Given the prior constructed here (based on life
310 history values via empirical M estimators) drove the estimation of M and profiling showed no
311 information to delineated M values >0.4, fixing M to 0.39 year⁻¹ seemed a very reasonable
312 decision when determining a reference model.

313 This model also showed sensitivity to steepness for several model derived quantities, a
314 common result as changes in the steepness value usually causes major uncertainty in the

315 estimation of management quantities (Zhou, 2007). But the model also was unable to estimate
316 steepness given the lack of strong contrast in population biomass and recruitment, despite the
317 u-shaped population dynamics in the model using CPUE (Lee, Maunder, Piner, & Methot,
318 2012; Magnusson & Hilborn, 2007).

319 Length composition data is one of the easier data sources to collect for many species, though
320 non-representative sampling can potentially cause bias in interpreting sampled lengths. Length
321 data are a central component for age-structured models, especially when aging data are
322 typically not available (Heery & Berkson, 2009), providing information on gear selectivity,
323 recruitment pulses, and stock status (as well as life history parameter information in some
324 situations). Length data can suffer from systematic errors during the sampling of catch that
325 make it unrepresentative of the true catch. While the causes of bias in sampling fishery-
326 dependent length composition data are recognized (e.g, non-random sample collection, limited
327 access to fishery catch or poor sampling design), the effect of it on stock assessment is always
328 not straightforward (Gerritsen & McGrath, 2007; Heery & Berkson, 2009). Here we
329 demonstrate that for *L. purpureus*, correcting for the main source of sampling error (no
330 sampling of the exported portion of catch) had little effect on model outputs. While this lack of
331 model sensitivity points to model robustness to this particular data scenario,
332 representativeness in the length data should always either be ensured through proper
333 sampling design or evaluated in the model with the exploration of data scenarios.

334 One of the biggest sources of uncertainty in the red snapper model was the inclusion of the
335 CPUE-based abundance index. CPUE misspecification can cause a significant weakness in
336 the model performance when linking the population trend to the abundance index (Methot Jr
337 & Wetzel, 2013; Wiedenmann & Jensen, 2017). In the case of this model, the final trend in the
338 population dynamics demonstrated a major dichotomy in results depending on the treatment
339 of the CPUE index. Removing it caused the population to continue to decline (as the catch
340 continued to increase) instead of rebound. This also demonstrates how the signal in the index
341 was different to that of the length composition data. Competing signals in data sources are

342 very common in integrated stock assessments, and must be resolved using data weighting or,
343 ideally, alternative model specification (Maunder and Piner, 2017). Such data weighting
344 choices are a major consideration when model building and defining appropriate sensitivity
345 analyses.

346 Fishery-dependent CPUE is known to vary over time violating the assumption of being
347 proportional to abundance. Several methods have been employed to incorporate time-varying
348 catchability into stock assessments (e. g. random walk) but fishery-dependent CPUE generally
349 need standardization (Wilberg, Thorson, Linton, & Berkson, 2009). CPUE data series can be
350 standardized to account for a variety of factors, however standardization can only correct for
351 measured factors and require available data for each factor (Wilberg et al., 2009). Information
352 on stock spatial and temporal variability and technological changes in the French Guiana red
353 snapper fishery is fragmentary and incomplete, thus, CPUE standardisation for this red
354 snapper model is currently not possible. The incorporation of a non-linear power function
355 between the CPUE index and the catchability suggests possible hyperdepletion that added
356 more uncertainty to the model outputs. One possible mechanism explaining this result is that
357 the grouping behaviours of snappers can lead to localized depletions as already showed for
358 *Lutjanus* spp. In the Gulf of Mexico (Saul, Brooks, & Die, 2020). Nevertheless, those results
359 should be interpreted with caution since the CPUE integrated in the model were not
360 standardized. Future use of CPUE as an index in this assessment should consider the
361 possibility of incorporating time-varying or non-proportional catchability if additional data could
362 be collected to improve CPUE standardization and application.

363 The VPA approach has a long history of application to major fished stocks in French Guiana
364 and several European regions but this methods requires a relatively complete data set and
365 can often accommodate only the most recent period of the fishery since age composition data
366 are rarely available for the beginning of the fishery (Stewart & Martell, 2015). The VPA
367 approach requires a complete catch-at-age time series that is often estimated from cohort
368 slicing of length data (Ailloud et al., 2015). This type of procedures can introduce a large and

369 unpredictable uncertainty that can influence VPA assessments (Carruthers, Kell, & Palma,
370 2017). The SS model does not need catches-at age and can directly integrate length datasets
371 using growth parameters and uncertainty in length at age, thus integrating this uncertainty in
372 the assessment (Methot Jr & Wetzell, 2013). SS and VPA follow a similar process but in
373 opposite directions (VPA is a backward projection model while SS is a forward projection
374 model) and adopt a different selectivity approach (Punt, Hurtado-Ferro, & Whitten, 2014) and
375 treatment of recruitment. In the recent years, French Guiana stock assessments has been
376 performed with both methods to compare the results. The possibility to use SS exclusively for
377 future assessments is now a consideration. The differences result in a notable difference in
378 absolute biomass estimation among the approaches. Stewart and Martell (2015) also saw a
379 biomass difference comparing VPA and SS models. The catch and length version of the SS
380 model, which is the most similar SS model to the data used in the VPA model, was also
381 different in both population trend and biomass size. These models all give very different
382 measures of absolute and/or relative stocks size, and thus management should consider the
383 most appropriate way to weight these different model specifications in order to inform
384 management (Stewart & Martell, 2015). The SS model allows for flexibility and uncertainty
385 specification and should be preferred over VPA for further management scenarios.

386 **Conclusion**

387 Dealing with data-limited fisheries and unreported times series can be particularly challenging,
388 and model misspecification and data treatments can cause important changes in the model
389 outputs and management suggestions. Prioritizing some of the data and down-weighting
390 others can be a solution to reduce conflicts but it can be difficult to choosing these weightings
391 (Ichinokawa, Okamura, & Takeuchi, 2014). These conflicts instead should be confronted with
392 model exploration to avoid model misspecification (Wang & Maunder, 2017) or re-evaluation
393 of the representativeness of the data in question. Thorough sensitivity analysis and even
394 simulation analysis may be needed to identify potential bias and misspecifications.

395 Our results showed that for French Guiana *L. purpureus* the SS model provided a flexible and
396 viable method to assess the exploitation status of the stock and the uncertainty to model
397 specification and data set choices given the limitations in available data and life history inputs.
398 The available data were insufficient for the estimation of natural mortality and steepness,
399 necessitating a sensitivity exploration through a likelihood profile to understand how model
400 outputs were affected by the values of these parameters. This model also showed sensitivity
401 to data inputs, as the CPUE index seems to contrast with the length composition data-set.
402 Whether this is due to truly different signals in the data, lack of proper standardization in the
403 CPUE, or unrepresentativeness of the data is not known at this time, but it does pinpoint a
404 critical research topic to improve future stock assessments of red snapper. Despite the above
405 uncertainties, all models were depicting a similar trend with notable stock depletion in the late
406 90s. Biomass is recovering in recent years when using the CPUE abundance index (~60% of
407 the unfished spawning biomass) despite stable fishing mortality. To preserve this important
408 economic resource, new data collections (e.g. measuring lengths of all catches; improving the
409 quality of the CPUE time series with electronic monitoring; development of a fishery-
410 independent survey; collecting ageing structures) can be added directly to the SS model
411 configured here, while enforceable management measures (e.g. hook size regulations;
412 developing spatial and temporal restrictions; limit illegal activities; catch limits) should be
413 explored.

414 **Acknowledgments**

415 We would like to tanks all the SIH team (E. Mansuy and all the VSC) for collecting the data.

416 We would also like to acknowledge the fisherman committee (CPMEM) and all fishermen for

417 their willingness to share information with us. This work has been initiated in the the

418 IFREMER/WECAFC/CRFM Working Group on Shrimp and Groundfish of the North Brazil-

419 Guiana Shelf allowing, funding, the collaboration with NOAA (J. Cope).

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585 **Tables**586 **Table 1: Life history parameters employed in this study**

von Bertalanffy growth coefficient (k)	0.12 year ⁻¹	(Rivot et al., 2000)
von Bertalanffy asymptotic length (L _{inf})	105 cm	(Rivot et al., 2000)
length-weight allometric parameter (b)	2.95455	(Lampert, Achoun, & Levrel, 2013)
length-weight scaling parameter (a)	1.97E-05	(Lampert et al., 2013)
maximum age	13 year	(Rivot et al., 2000)
maximum length	88 cm	(Rivot et al., 2000)
Length at 50% maturity	32 cm	

587

588 **Figure legends**

589 Fig. 1: Catches and catch per unit effort (CPUE) data employed in red snapper SS model.

590 Fig. 2: Comparison of the main model outputs for the reference SS model, the SS model
591 without catch per unit effort (CPUE) and the virtual population analysis (VPA).

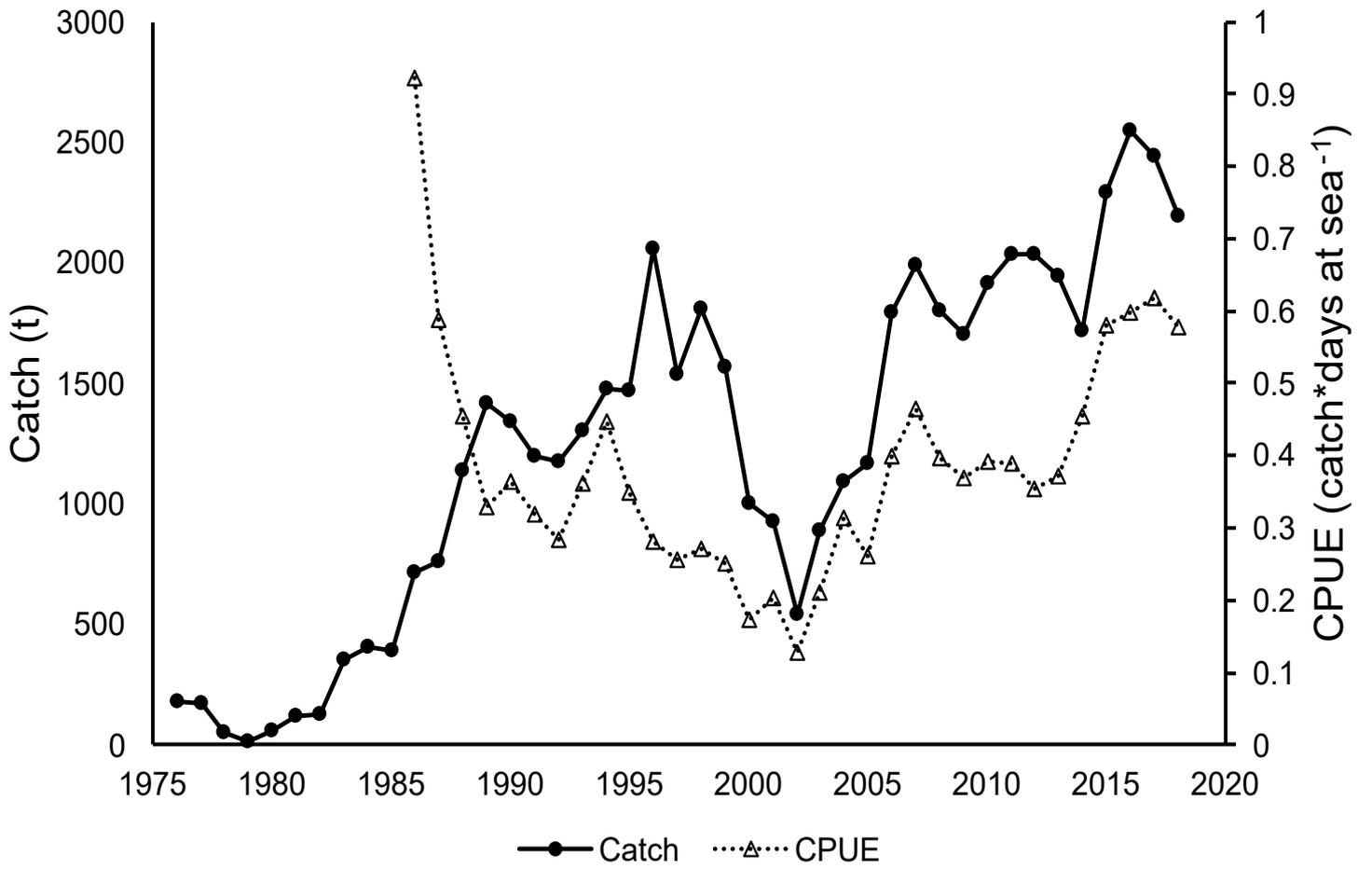
592 Fig. 3: Likelihood profile for natural mortality and derived quantities (initial spawning output
593 (SO_0); spawning output in 2018 (SO_{2018}), stock status (SO_{2018}/SO_0) in the French Guiana red
594 snapper SS model. The natural mortality of 0.39 estimated by “The Natural Mortality tool” is
595 showed by a grey dot.

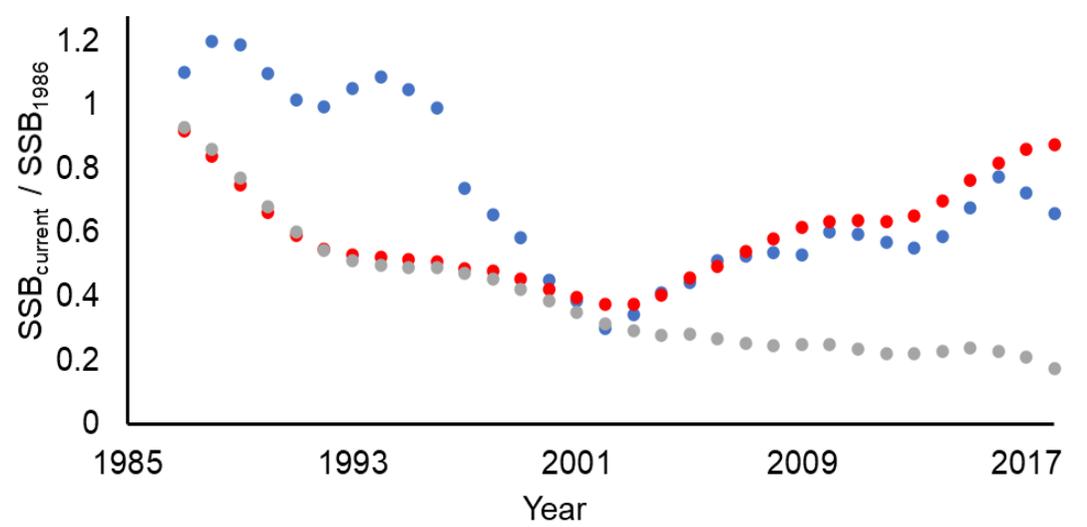
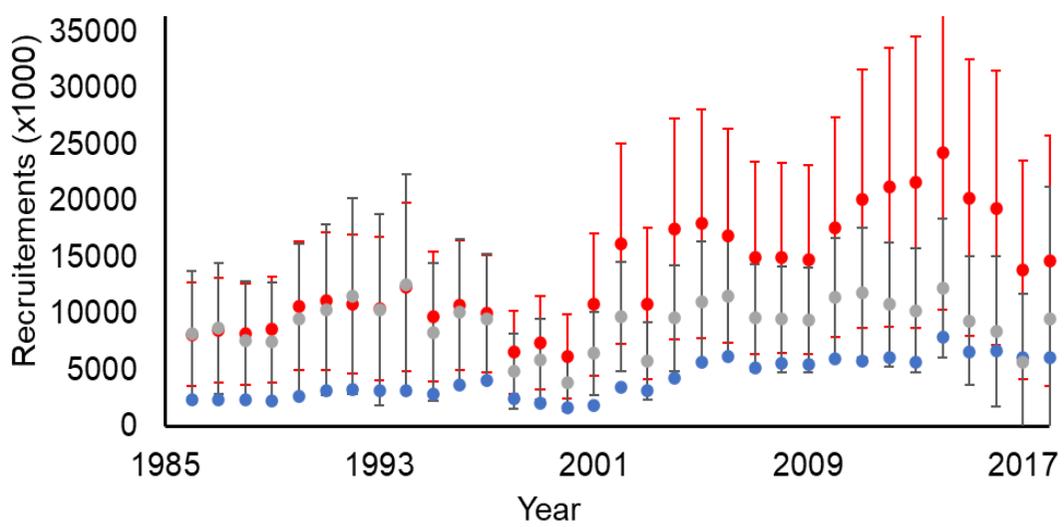
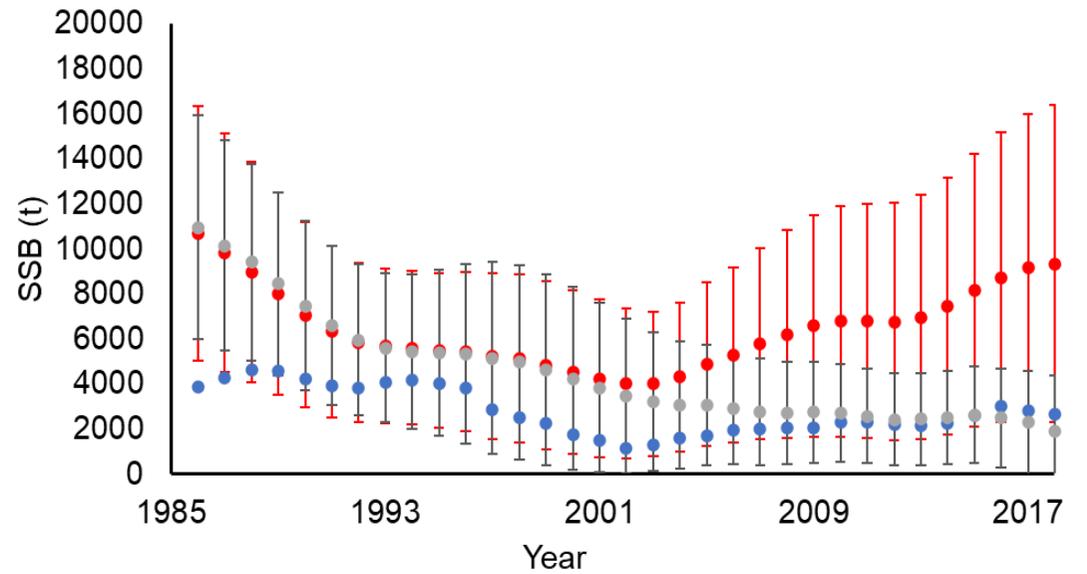
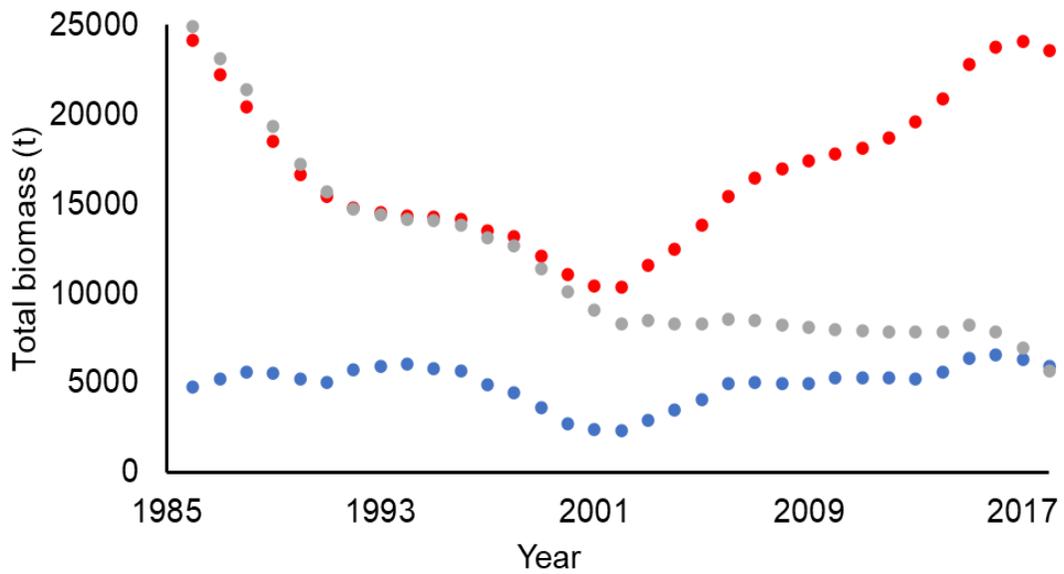
596 Fig. 4: Likelihood profile for steepness (h) and derived quantities (initial spawning output (SO_0);
597 spawning output in 2018 (SO_{2018}), stock status (SO_{2018}/SO_0) in the French Guiana red snapper
598 SS model. The steepness value of 0.7 estimated by Fishlife is showed by a grey dot.

599 Fig. 5: Comparison of the main model outputs and index fit for the reference model and the
600 model using the modified length composition dataset.

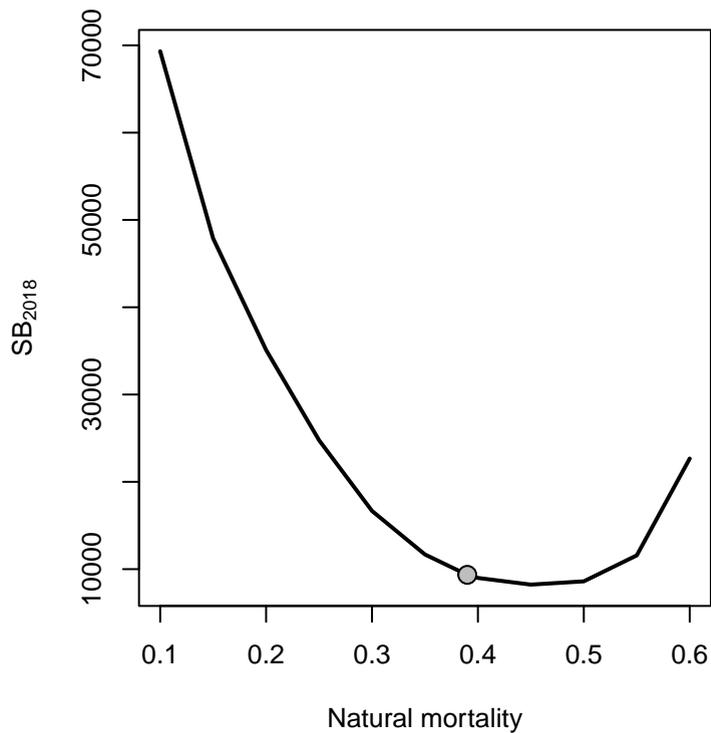
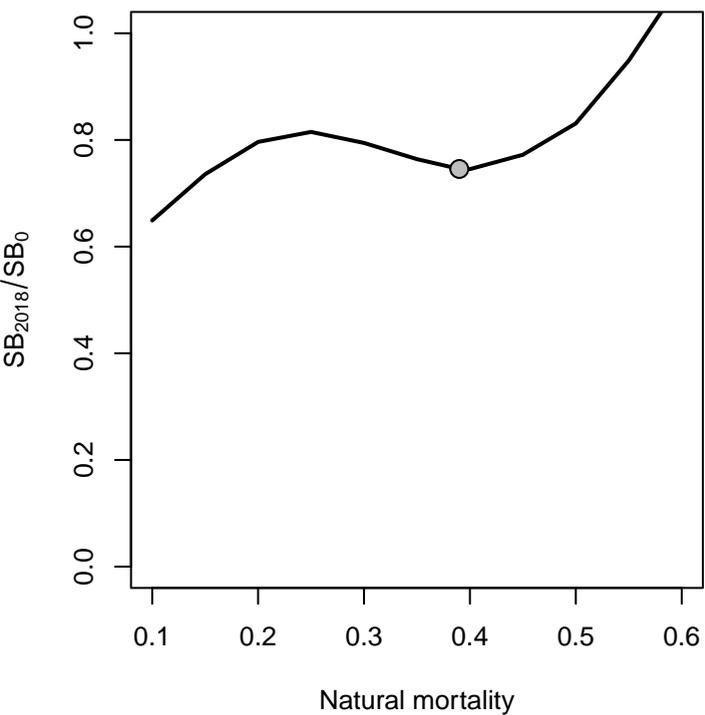
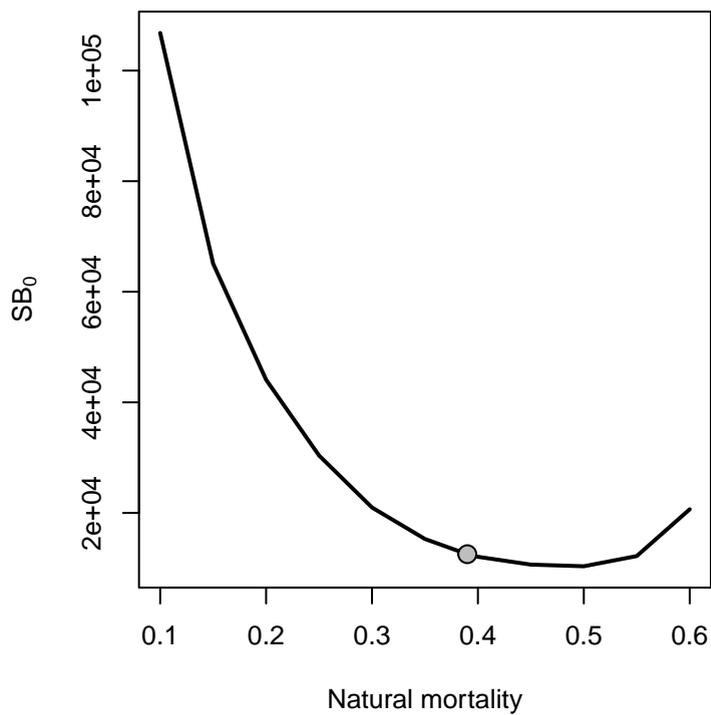
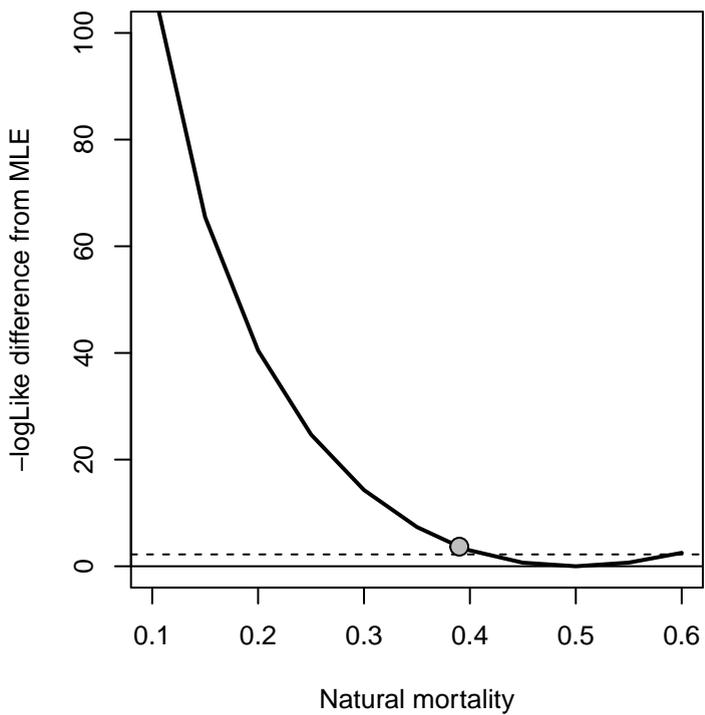
601 Fig. 6: Comparison of the main model outputs and index fit for the reference model and the
602 model without catch per unit effort (CPUE).

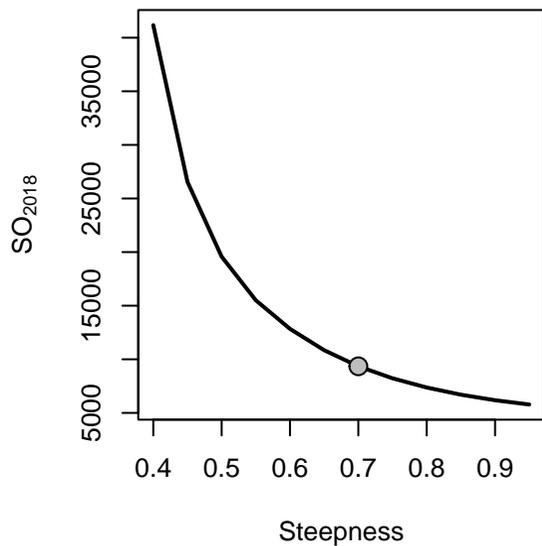
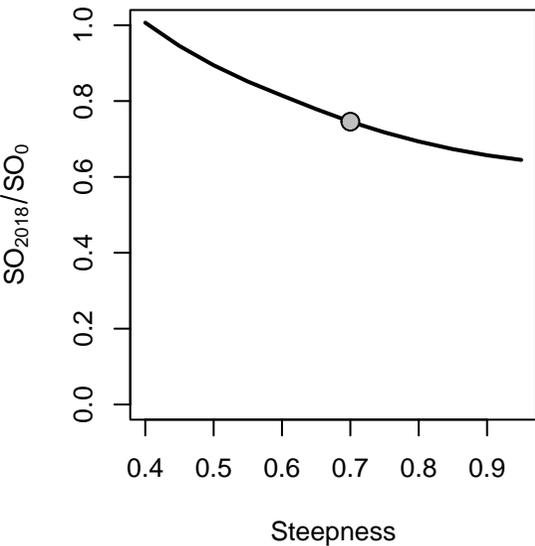
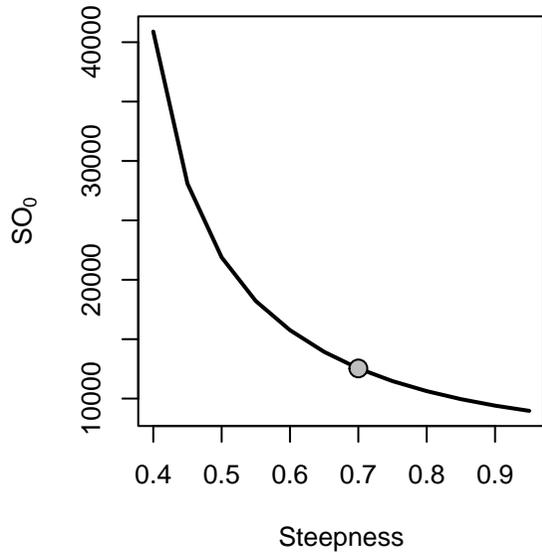
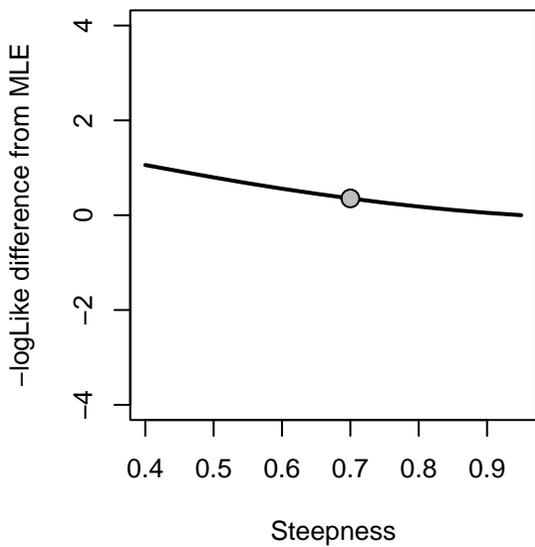
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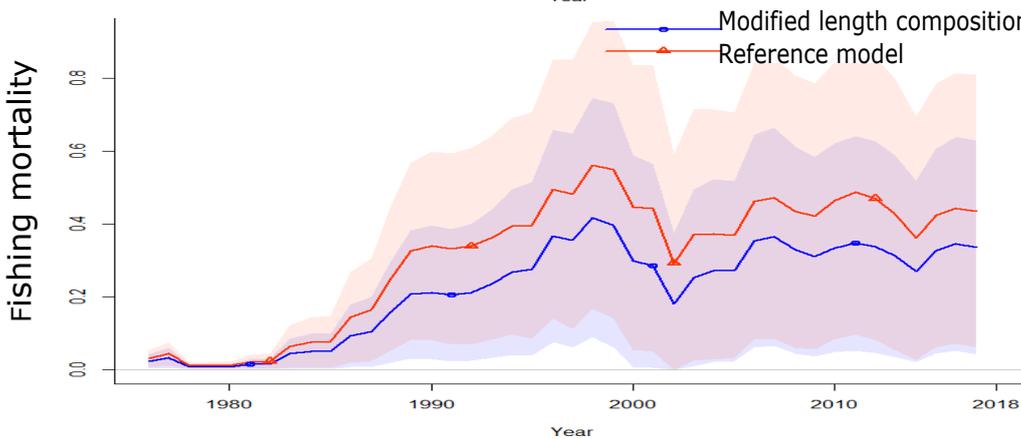
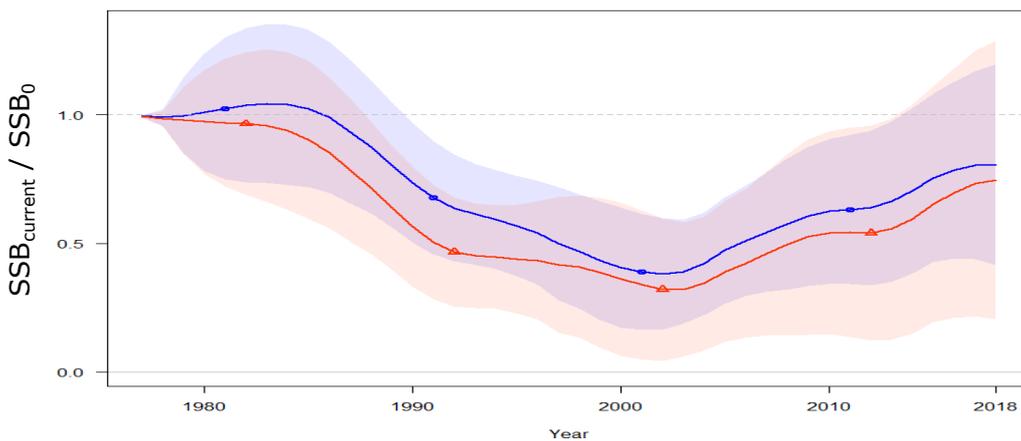
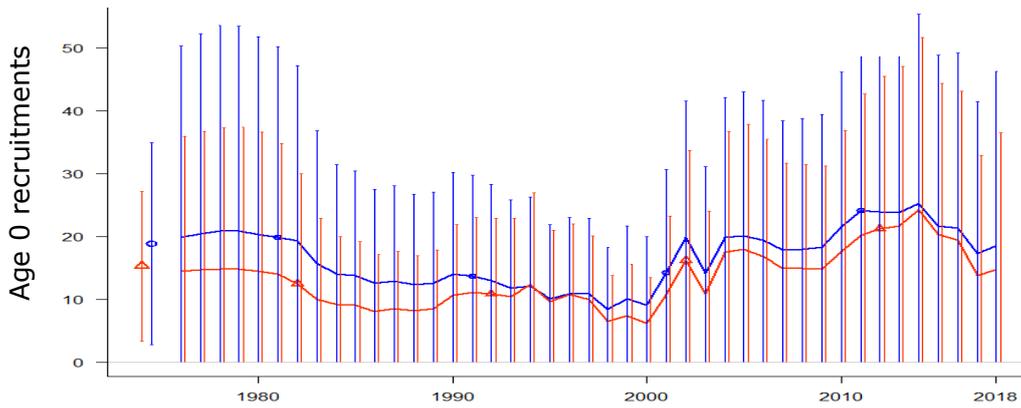
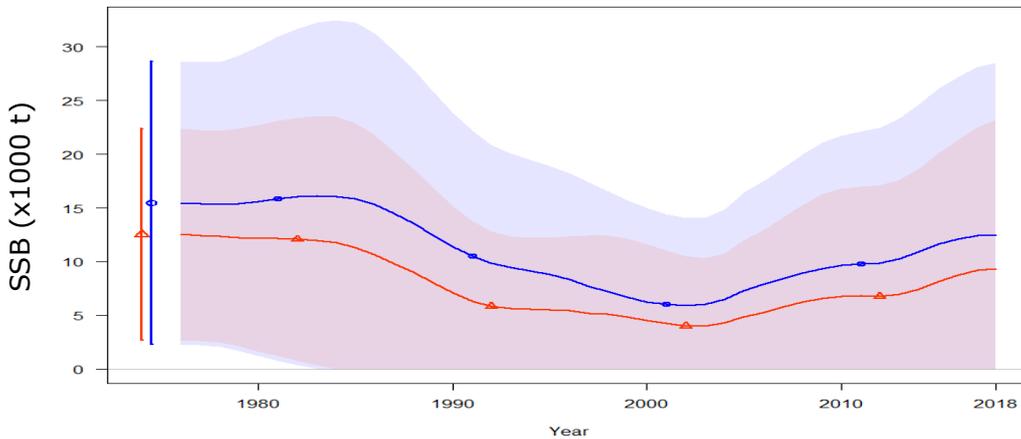
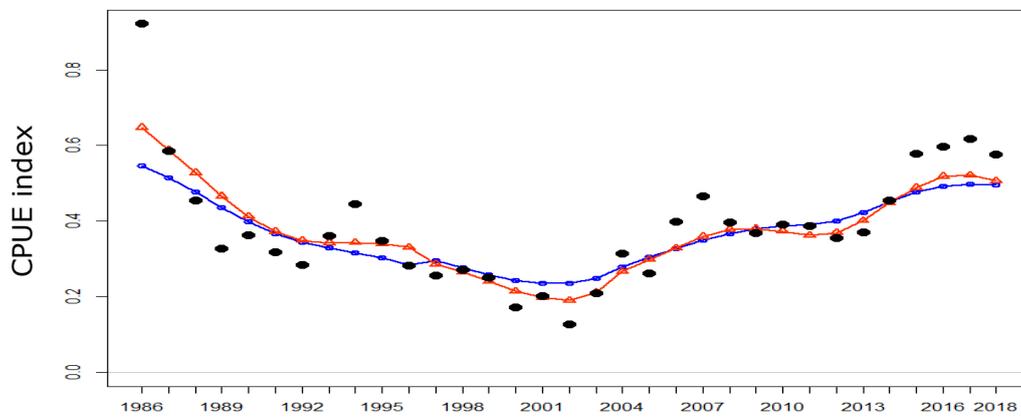


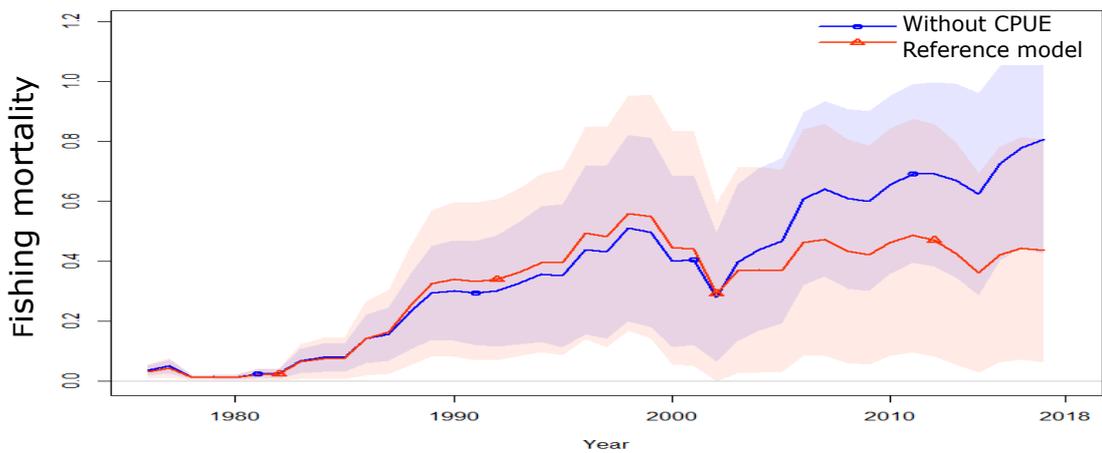
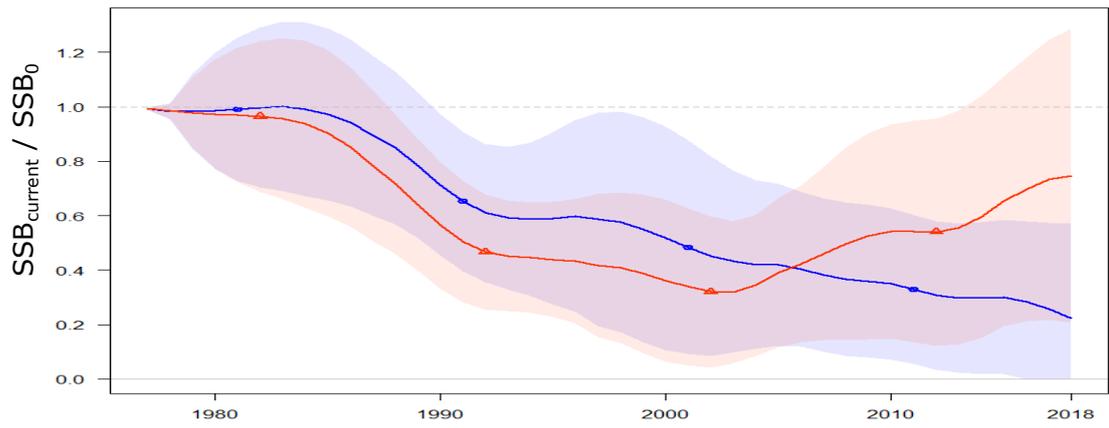
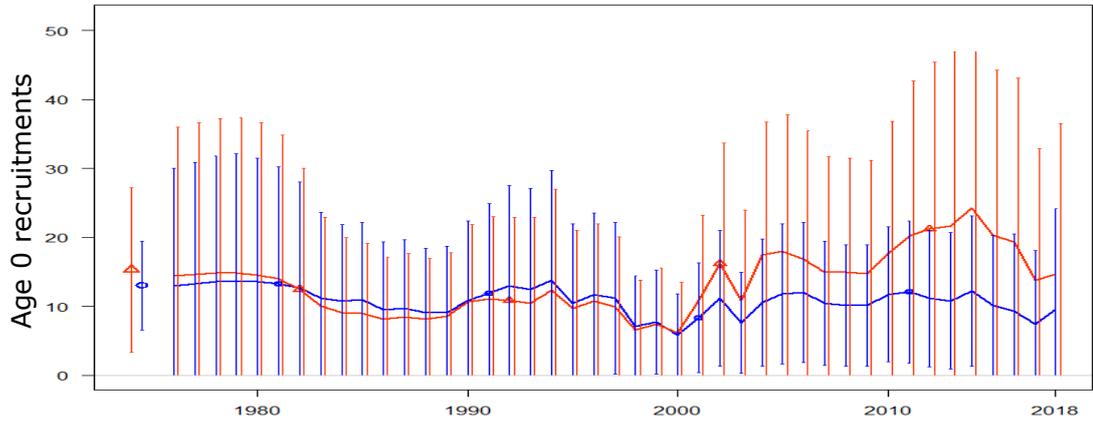
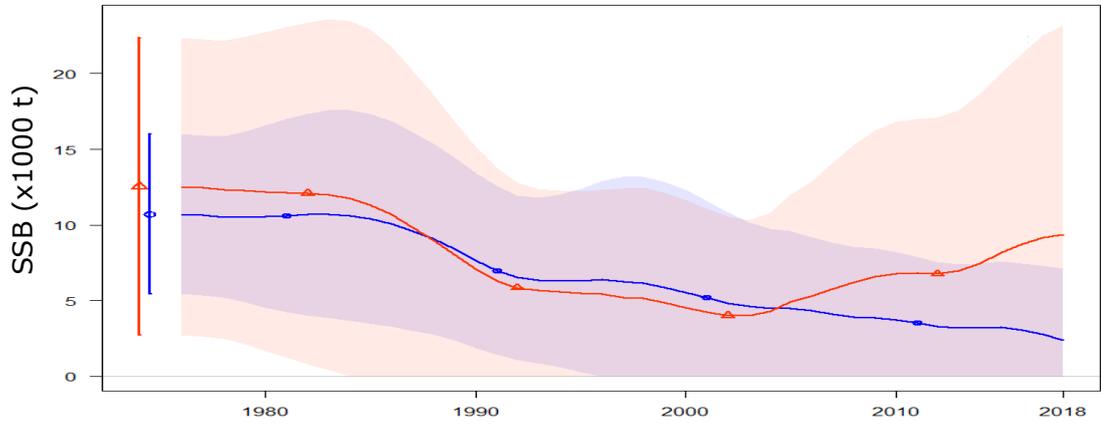
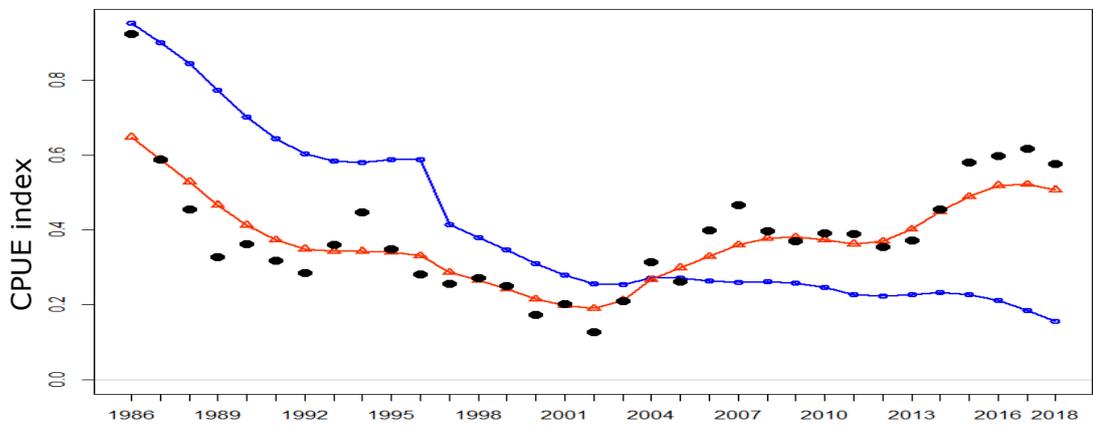


- VPA
- Reference SS model
- Without CPUE SS model









List of Appendix:

Appendix A

List of the parameters used in the reference model. 2

Appendix B

Diagnostics for the stock synthesis model without time-block selectivity of red snapper in French Guiana. 5

Appendix C

Diagnostics and likelihood profiles for the reference stock synthesis model of red snapper in French Guiana. 10

Appendix D

Diagnostics and likelihood profiles for the stock synthesis model without the catch per unit effort index of red snapper in French Guiana. 14

Appendix E

Comparison of the length composition distribution and selectivity. 19

Appendix F

Comparison of the reference model and the model using a power link function. 20

Appendix A

List of the parameters used in the reference model.

Table A1: List of parameters used in the base model, including estimated values and standard deviations (SD), bounds (minimum and maximum), estimation phase (negative values not estimated), status (indicates if parameters are near bounds), and prior type information (mean and SD).

Parameter	Value	Phase	Bounds	Status	SD	Prior (Exp.Val, SD)
NatM p 1 Fem GP 1	0.39	-5	(0.001, 0.6)	-	-	Log Norm (-0.63, 0.17)
L at Amin Fem GP 1	13	-3	(5, 25)	-	-	None
L at Amax Fem GP 1	85	-3	(30, 45)	-	-	None
VonBert K Fem GP 1	0.12	-3	(0.1, 0.4)	-	-	None
CV young Fem GP 1	0.1	-5	(0.03, 5)	-	-	None
CV old Fem GP 1	0.1	-5	(0.03, 5)	-	-	None
Wtlen 1 Fem GP 1	0	-99	(0, 3)	-	-	None
Wtlen 2 Fem GP 1	2.955	-99	(2, 4)	-	-	None
Mat50% Fem GP 1	32	-99	(20, 40)	-	-	None
Mat slope Fem GP 1	-1	-99	(-2, 4)	-	-	None
Eggs/kg inter Fem GP 1	1	-99	(0, 6)	-	-	None
Eggs/kg slope wt Fem GP 1	0	-99	(-3, 5)	-	-	None
NatM p 1 Mal GP 1	0.39	-5	(0.001, 0.6)	-	-	Log Norm (-0.63, 0.17)
L at Amin Mal GP 1	13	-2	(5, 25)	-	-	None
L at Amax Mal GP 1	85	-2	(30, 45)	-	-	None
VonBert K Mal GP 1	0.12	-3	(-1, 1)	-	-	None
CV young Mal GP 1	0.1	-5	(-5, 5)	-	-	None
CV old Mal GP 1	0.1	-5	(-5, 5)	-	-	None
Wtlen 1 Mal GP 1	0	-99	(0, 3)	-	-	None
Wtlen 2 Mal GP 1	2.955	-99	(2, 4)	-	-	None
CohortGrowDev	1	-99	(0, 2)	-	-	None
FracFemale GP 1	0.5	-99	(0.01, 0.99)	-	-	None
SR LN(R0)	9.633	1	(3, 31)	OK	0.4	None
SR BH steep	0.7	-4	(0.2, 1)	-	-	Sym Beta (0.7, 0.05)
SR sigmaR	0.6	-4	(0, 2)	-	-	None
SR regime	0	-4	(-5, 5)	-	-	None
SR autocorr	0	-99	(0, 0)	-	-	None
Main RecrDev 1976	0.014	1	(-6, 6)	act	0.593	dev (NA, NA)
Main RecrDev 1977	0.031	1	(-6, 6)	act	0.596	dev (NA, NA)
Main RecrDev 1978	0.046	1	(-6, 6)	act	0.598	dev (NA, NA)
Main RecrDev 1979	0.047	1	(-6, 6)	act	0.596	dev (NA, NA)
Main RecrDev 1980	0.031	1	(-6, 6)	act	0.588	dev (NA, NA)
Main RecrDev 1981	-0.004	1	(-6, 6)	act	0.563	dev (NA, NA)
Main RecrDev 1982	-0.116	1	(-6, 6)	act	0.513	dev (NA, NA)

Main RecrDev 1983	-0.33	1	(-6, 6)	act	0.455	dev (NA, NA)
Main RecrDev 1984	-0.417	1	(-6, 6)	act	0.41	dev (NA, NA)
Main RecrDev 1985	-0.422	1	(-6, 6)	act	0.388	dev (NA, NA)
Main RecrDev 1986	-0.519	1	(-6, 6)	act	0.393	dev (NA, NA)
Main RecrDev 1987	-0.459	1	(-6, 6)	act	0.384	dev (NA, NA)
Main RecrDev 1988	-0.487	1	(-6, 6)	act	0.39	dev (NA, NA)
Main RecrDev 1989	-0.418	1	(-6, 6)	act	0.396	dev (NA, NA)
Main RecrDev 1990	-0.177	1	(-6, 6)	act	0.387	dev (NA, NA)
Main RecrDev 1991	-0.113	1	(-6, 6)	act	0.392	dev (NA, NA)
Main RecrDev 1992	-0.118	1	(-6, 6)	act	0.39	dev (NA, NA)
Main RecrDev 1993	-0.148	1	(-6, 6)	act	0.412	dev (NA, NA)
Main RecrDev 1994	0.025	1	(-6, 6)	act	0.383	dev (NA, NA)
Main RecrDev 1995	-0.209	1	(-6, 6)	act	0.402	dev (NA, NA)
Main RecrDev 1996	-0.102	1	(-6, 6)	act	0.362	dev (NA, NA)
Main RecrDev 1997	-0.164	1	(-6, 6)	act	0.348	dev (NA, NA)
Main RecrDev 1998	-0.583	1	(-6, 6)	act	0.371	dev (NA, NA)
Main RecrDev 1999	-0.442	1	(-6, 6)	act	0.337	dev (NA, NA)
Main RecrDev 2000	-0.608	1	(-6, 6)	act	0.372	dev (NA, NA)
Main RecrDev 2001	-0.032	1	(-6, 6)	act	0.341	dev (NA, NA)
Main RecrDev 2002	0.386	1	(-6, 6)	act	0.319	dev (NA, NA)
Main RecrDev 2003	-0.014	1	(-6, 6)	act	0.406	dev (NA, NA)
Main RecrDev 2004	0.444	1	(-6, 6)	act	0.353	dev (NA, NA)
Main RecrDev 2005	0.44	1	(-6, 6)	act	0.366	dev (NA, NA)
Main RecrDev 2006	0.358	1	(-6, 6)	act	0.364	dev (NA, NA)
Main RecrDev 2007	0.219	1	(-6, 6)	act	0.375	dev (NA, NA)
Main RecrDev 2008	0.204	1	(-6, 6)	act	0.376	dev (NA, NA)
Main RecrDev 2009	0.183	1	(-6, 6)	act	0.38	dev (NA, NA)
Main RecrDev 2010	0.353	1	(-6, 6)	act	0.359	dev (NA, NA)
Main RecrDev 2011	0.488	1	(-6, 6)	act	0.368	dev (NA, NA)
Main RecrDev 2012	0.538	1	(-6, 6)	act	0.385	dev (NA, NA)
Main RecrDev 2013	0.552	1	(-6, 6)	act	0.398	dev (NA, NA)
Main RecrDev 2014	0.657	1	(-6, 6)	act	0.379	dev (NA, NA)
Main RecrDev 2015	0.461	1	(-6, 6)	act	0.409	dev (NA, NA)
Main RecrDev 2016	0.406	1	(-6, 6)	act	0.443	dev (NA, NA)
Main RecrDev 2017	0.002	1	(-6, 6)	act	0.523	dev (NA, NA)
Main RecrDev 2018	-0.005	1	(-6, 6)	act	0.593	dev (NA, NA)
ForeRecr 2019	0	5	(-6, 6)	act	0.6	dev (NA, NA)
LnQ base SURVEY1(2)	-9.847	-1	(-7, 5)	-	-	None
Q extraSD SURVEY1(2)	0	-4	(0, 0.5)	-	-	None
Size DbIN peak FISHERY(1)	40.729	2	(20, 80)	OK	2.232	None
Size DbIN top logit FISHERY(1)	-9.916	3	(-15, 3)	OK	78.08 9	None
Size DbIN ascend se FISHERY(1)	4.152	3	(-4, 12)	OK	0.376	None
Size DbIN descend se FISHERY(1)	3.398	3	(-10, 10)	OK	1.967	None

Size DbIN start logit FISHERY(1)	-999	-2	(-999, 15)	-	-	None
Size DbIN end logit FISHERY(1)	0.247	2	(-5, 15)	OK	0.682	None
SizeSel P1 SURVEY1(2)	-1	-4	(-5, 100)	-	-	None
SizeSel P2 SURVEY1(2)	-1	-5	(-5, 100)	-	-	None
Size DbIN peak FISHERY(1) BLK1repl 1997	31.27	4	(20, 80)	OK	1.006	None
Size DbIN top logit FISHERY(1) BLK1repl 1997	-4.547	5	(-15, 3)	OK	9.395	None
Size DbIN ascend se FISHERY(1) BLK1repl 1997	3.206	5	(-4, 12)	OK	0.314	None
Size DbIN descend se FISHERY(1) BLK1repl 1997	4.869	5	(-10, 6)	OK	1.014	None
Size DbIN end logit FISHERY(1) BLK1repl 1997	-1.464	4	(-5, 15)	OK	0.845	None

Appendix B

Diagnostics for the stock synthesis model without time-block selectivity of red snapper in French Guiana.

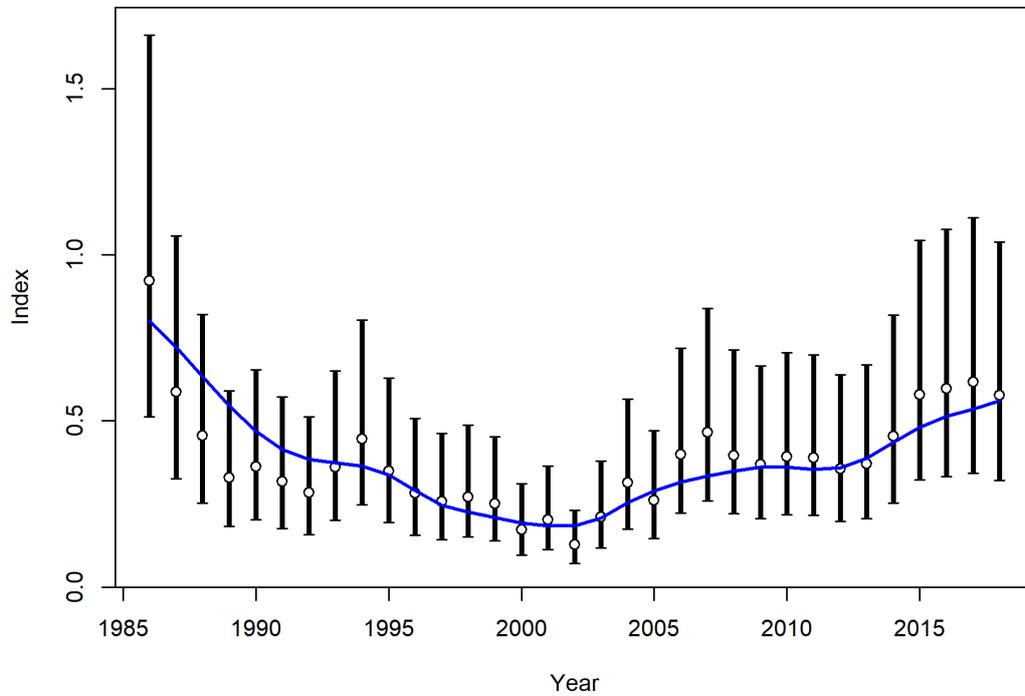


Figure B1: Fit of the model without time-block selectivity (blue line) to observed CPUE (mean and 95% uncertainty interval)

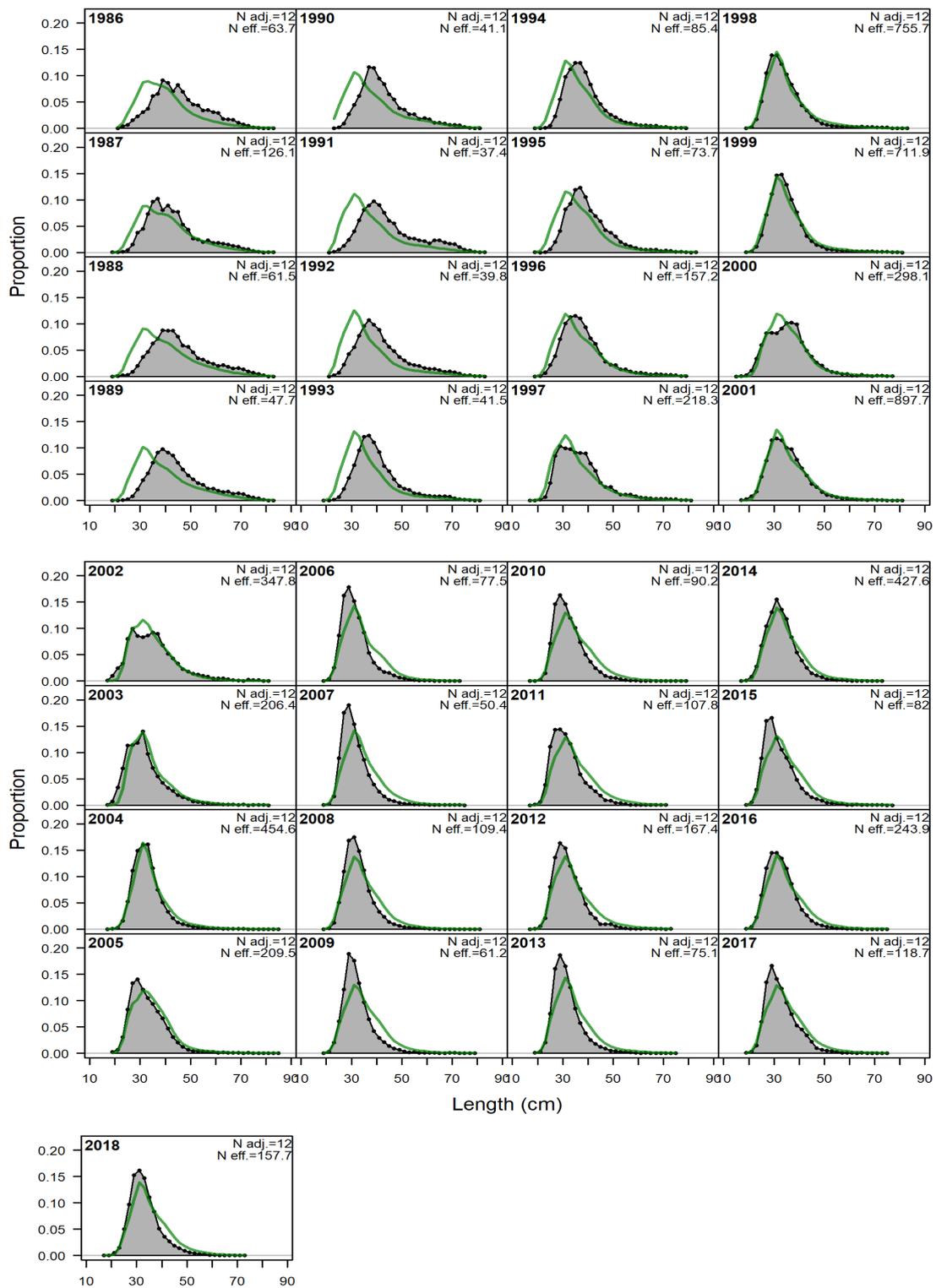


Figure B2: Fit of the model without time-block selectivity (green line) to observations on length compositions (grey area)

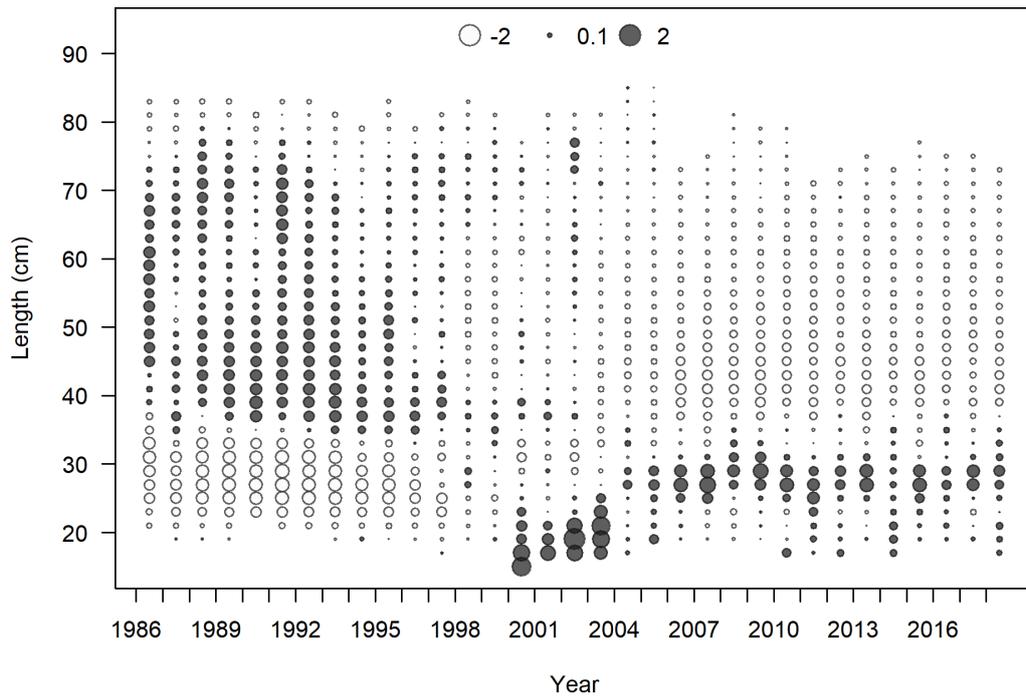


Figure B3: Pearson residuals of the model without time-block selectivity for length (closed bubbles are positive residuals and open bubbles are negative residuals)

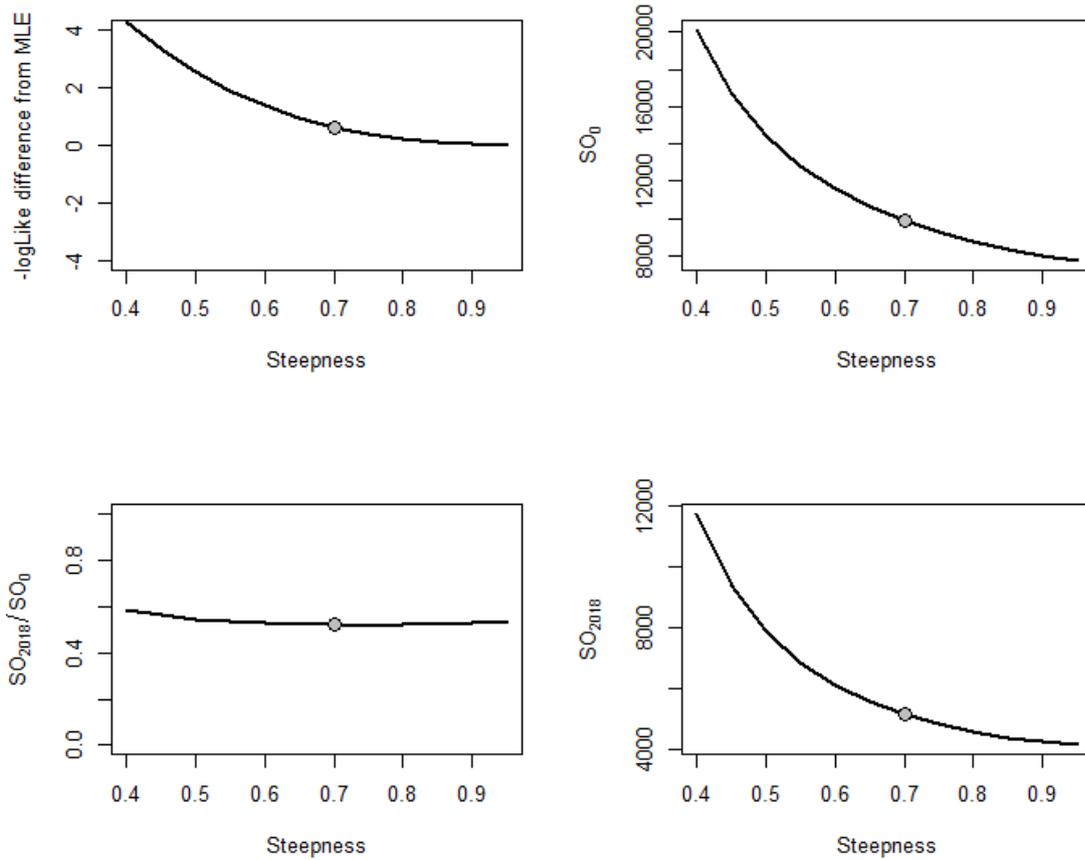


Figure B4 Likelihood profile in French Guiana red snapper SS model without time-block selectivity for steepness (h) and derived quantities (initial spawning output (SO_0); spawning output in 2018 (SO_{2018}), stock status (SO_{2018}/SO_0). The steepness value of 0.7 estimated by Fishlife is showed by a grey dot.

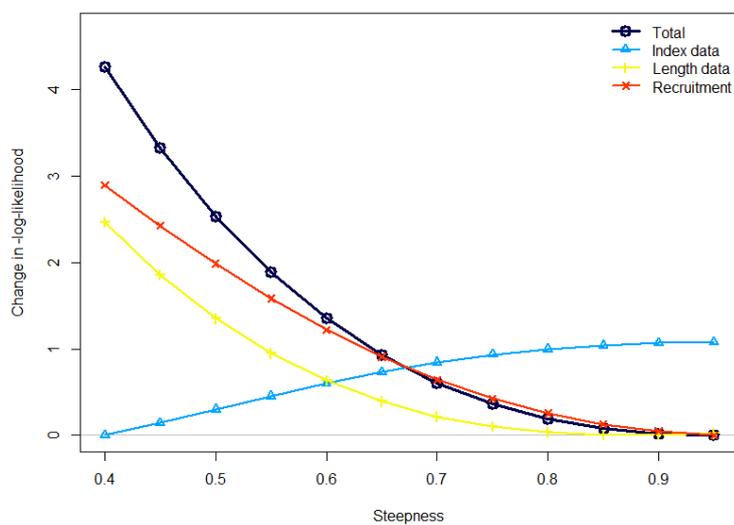


Figure B5 Likelihood component contributions for steepness profile in French Guiana red snapper SS model without time-block selectivity.

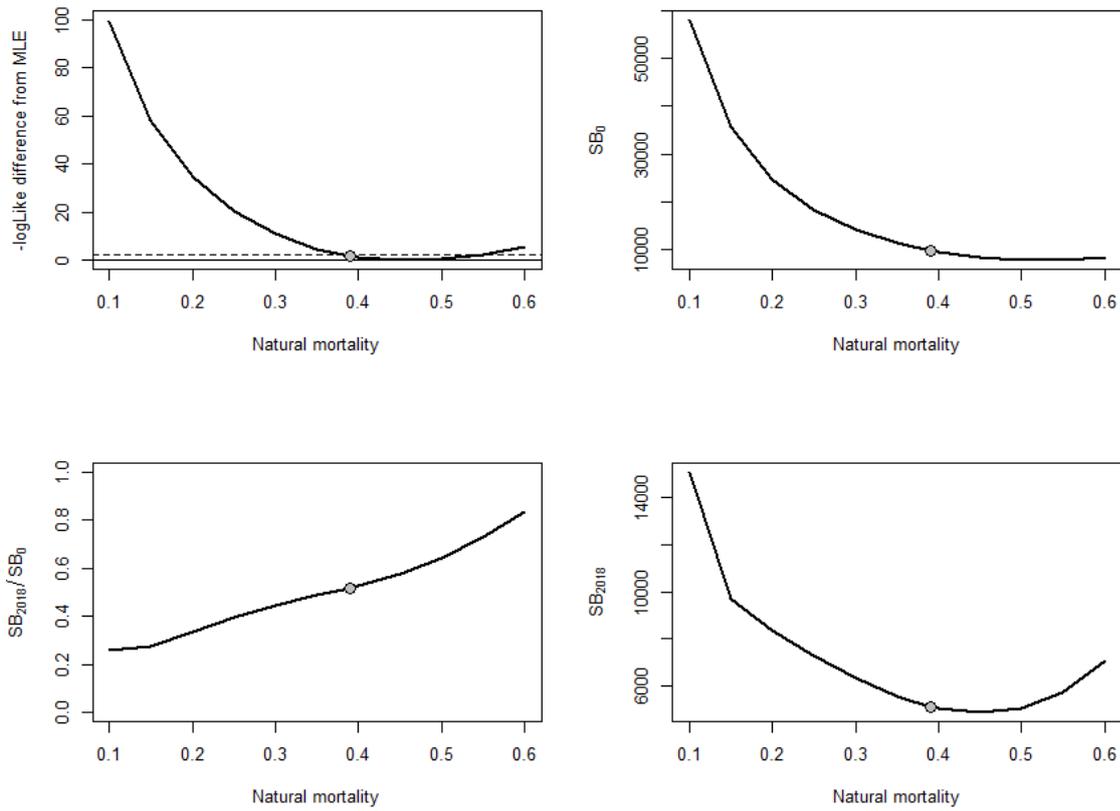


Figure B6 Likelihood profile in French Guiana red snapper SS model without time-block selectivity for natural mortality (M) and derived quantities (initial spawning output (SO_0); spawning output in 2018 (SO_{2018}), stock status (SO_{2018}/SO_0). The natural mortality of 0.39 estimated by “The Natural Mortality tool” is shown by a grey dot.

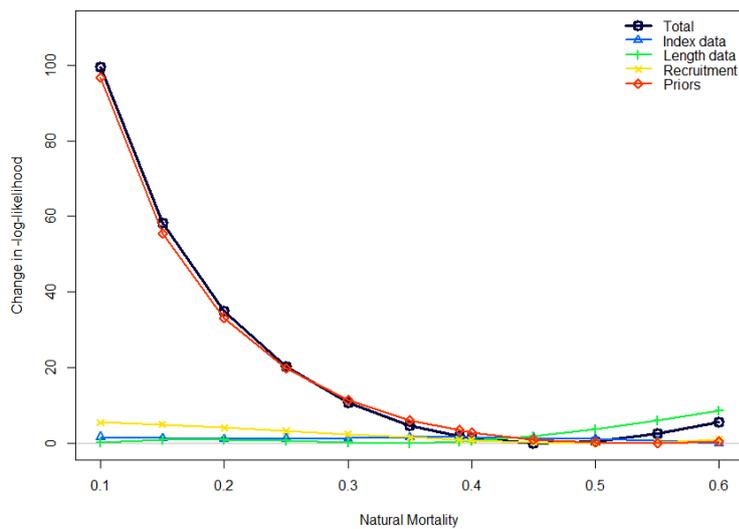


Figure B7 Likelihood component contributions for natural mortality (M) profile in French Guiana red snapper SS model without time-block selectivity.

Appendix C

Diagnostics and likelihood profiles for the reference stock synthesis model of red snapper in French Guiana.

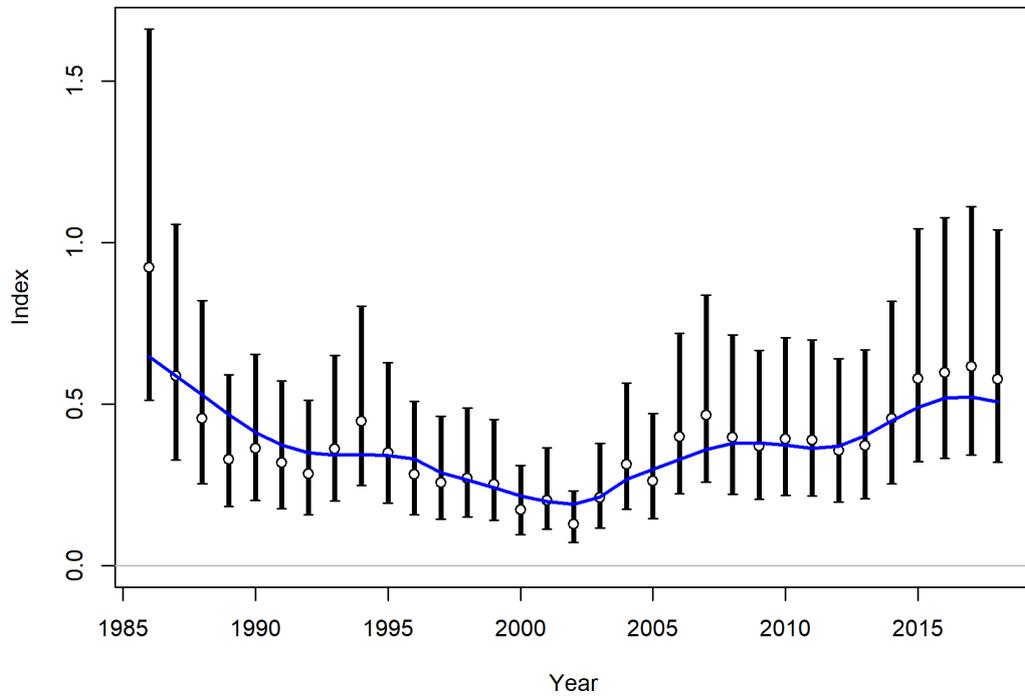


Figure C1: Fit of French Guiana red snapper SS reference model (blue line) to observed CPUE (mean and 95% uncertainty interval)

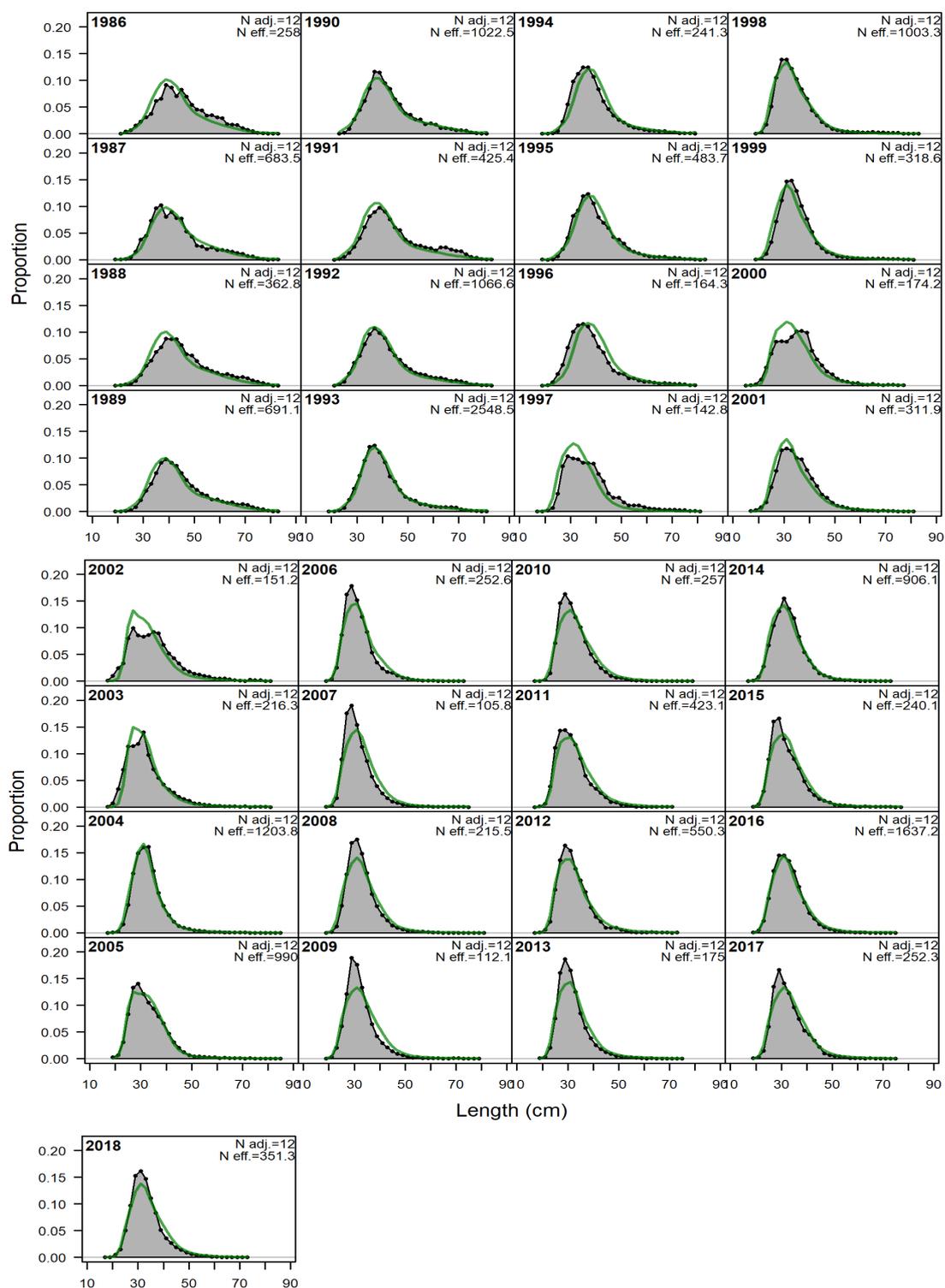


Figure C2: Fit of French Guiana red snapper SS reference model (green line) to observations on length compositions (grey area)

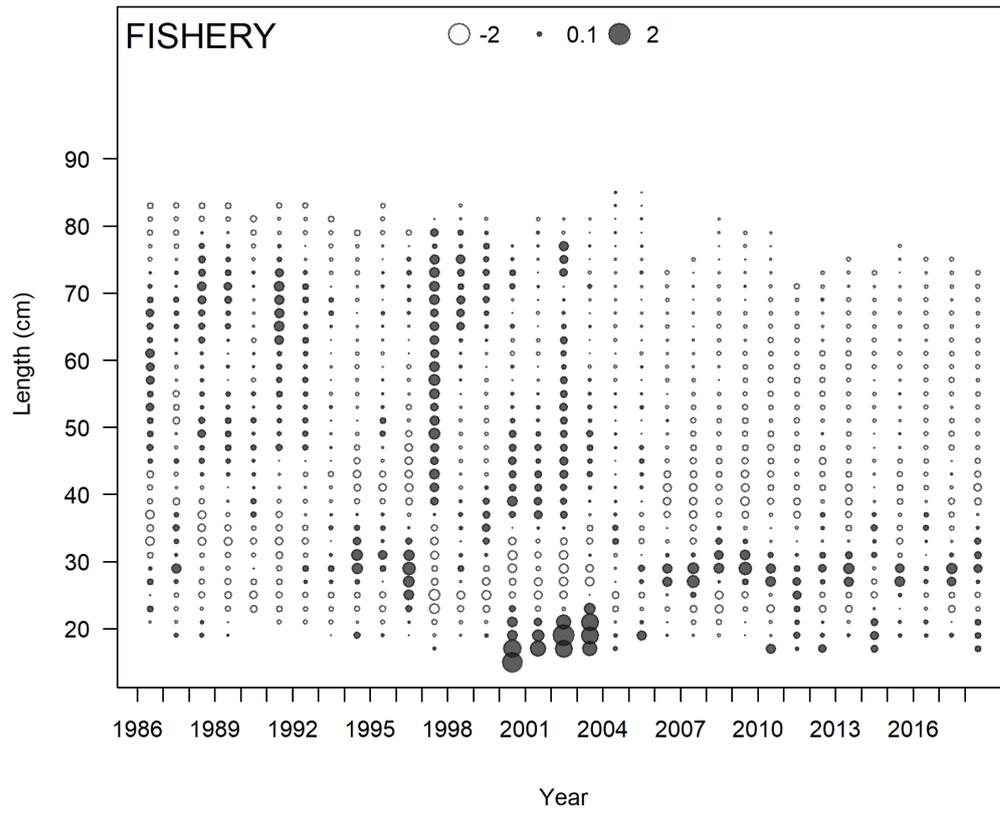


Figure C3: Pearson residuals of French Guiana red snapper SS reference model for length (closed bubbles are positive residuals and open bubbles are negative residuals)

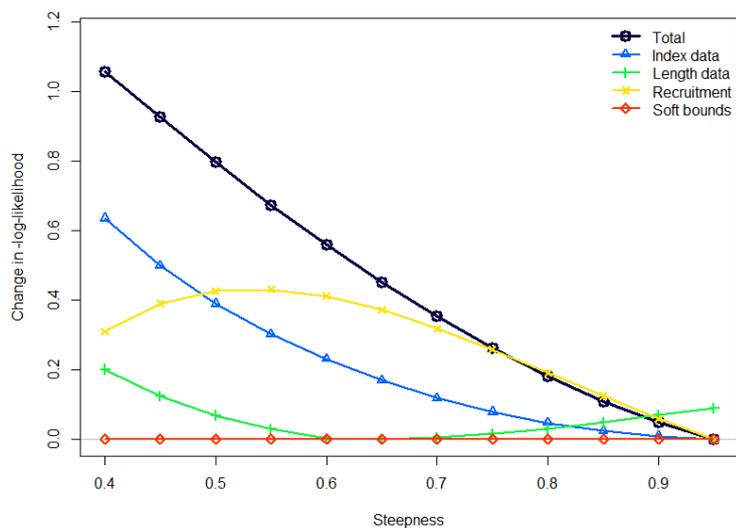


Figure C4: Likelihood component contributions for steepness profile in French Guiana red snapper SS reference model.

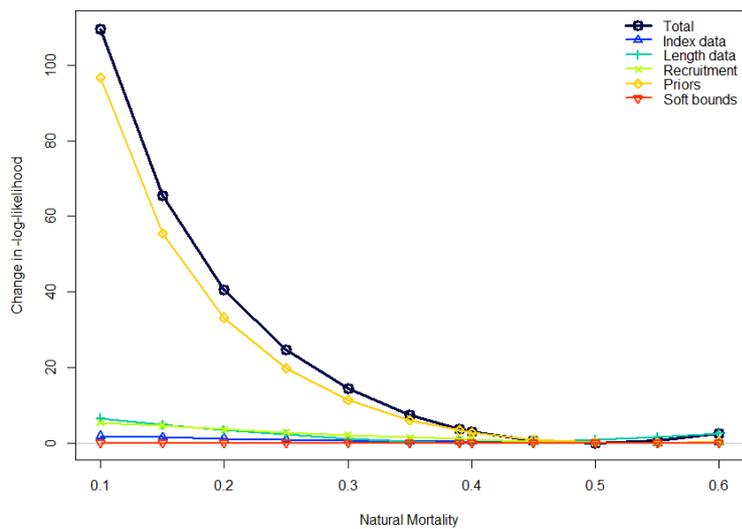


Figure C4: Likelihood component contributions for natural mortality (M) profile in French Guiana red snapper SS reference model.

Appendix D

Diagnostics and likelihood profiles for the stock synthesis model without the catch per unit effort index of red snapper in French Guiana.

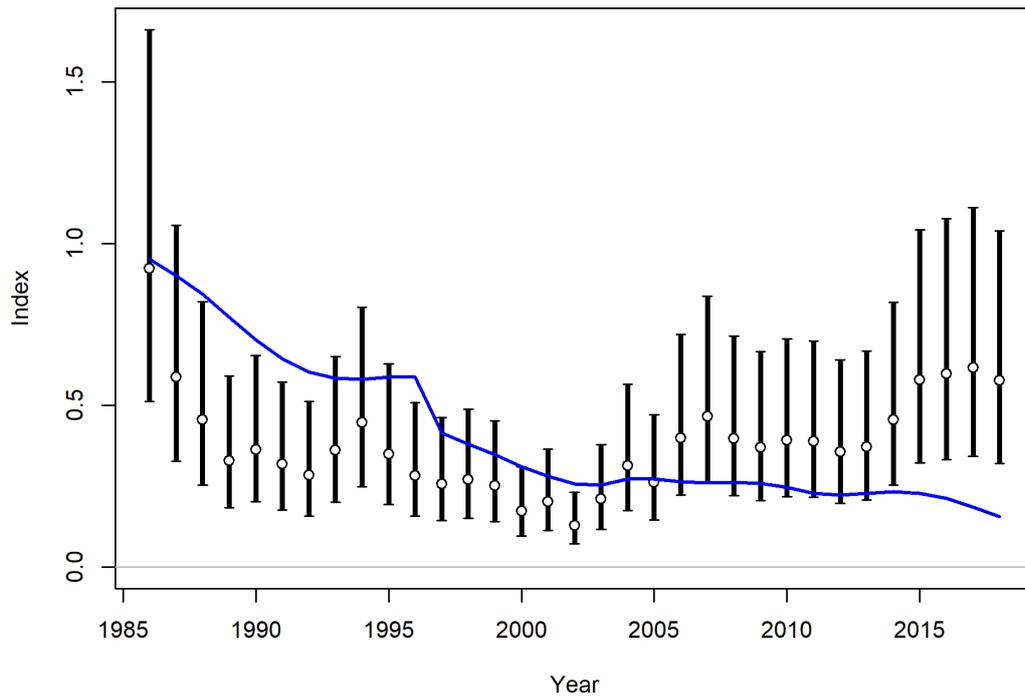


Figure D1: Fit of French Guiana red snapper SS model without the catch per unit effort index (blue line) to observed CPUE (mean and 95% uncertainty interval)

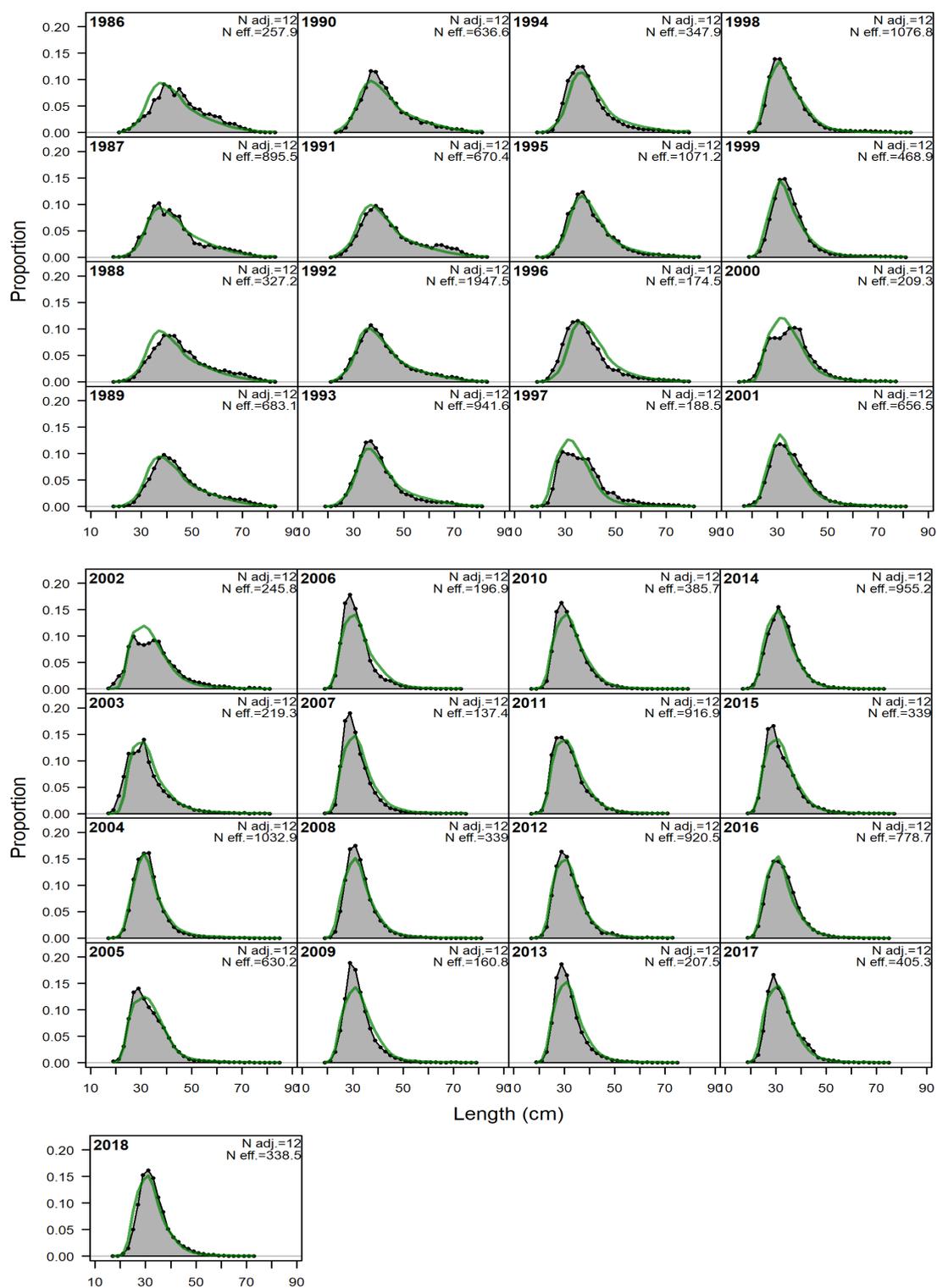


Figure D2: Fit of French Guiana red snapper SS without the catch per unit effort index (green line) to observations on length compositions (grey area)

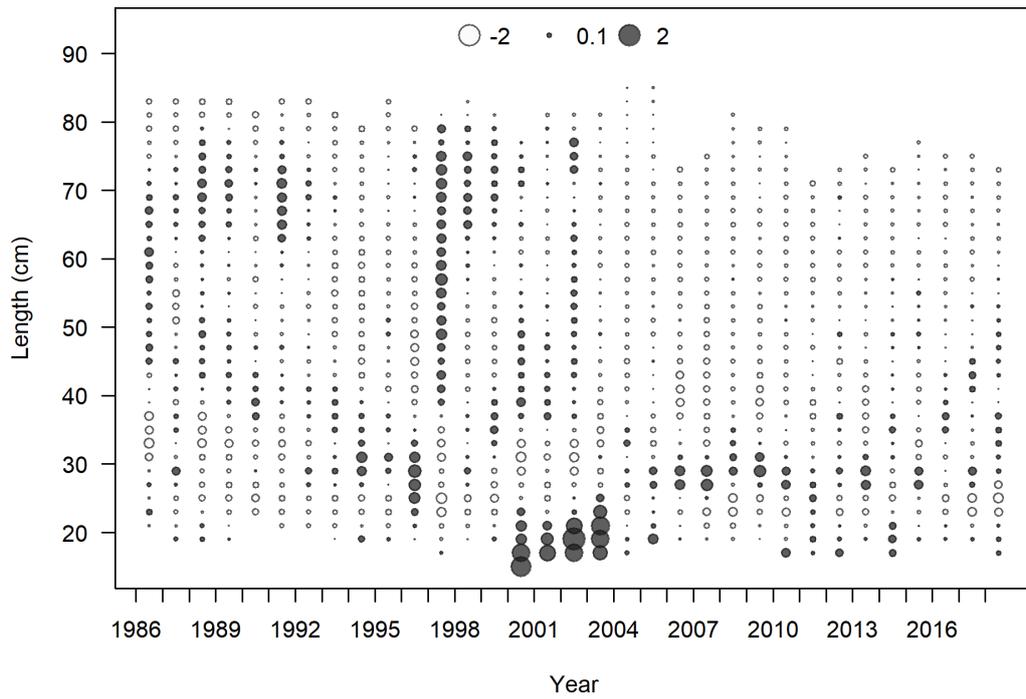


Figure D3: Pearson residuals of French Guiana red snapper SS without the catch per unit effort index for length (closed bubbles are positive residuals and open bubbles are negative residuals)

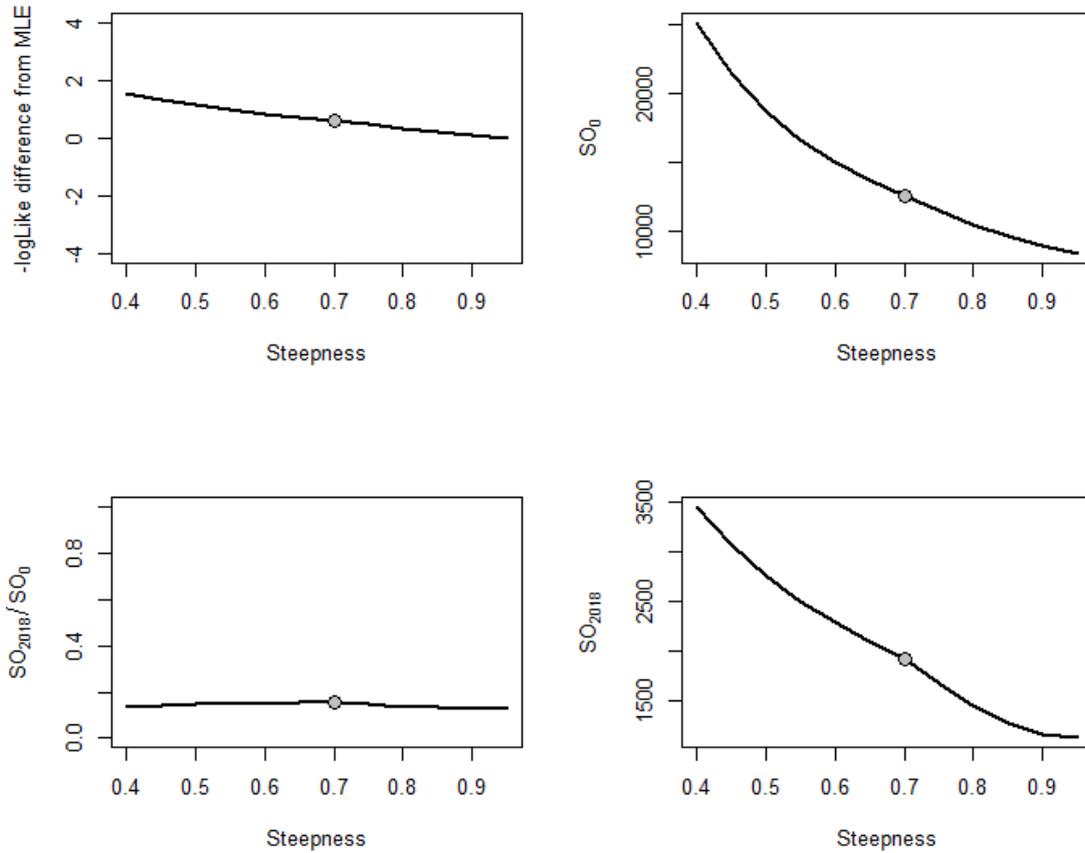


Figure D4: Likelihood profile in French Guiana red snapper SS model without the catch per unit effort index for steepness (h) and derived quantities (initial spawning output (SO_0); spawning output in 2018 (SO_{2018}), stock status (SO_{2018}/SO_0). The steepness value of 0.7 estimated by Fishlife is shown by a grey dot.

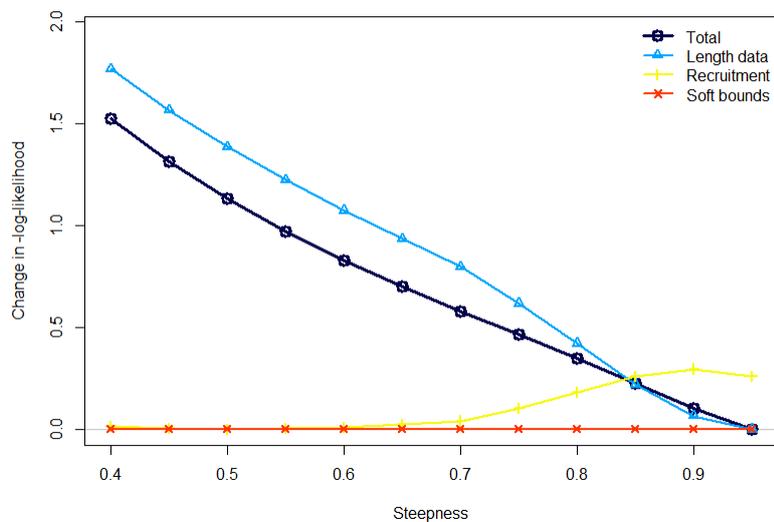


Figure D5: Likelihood component contributions for steepness profile in French Guiana red snapper SS model without the catch per unit effort index.

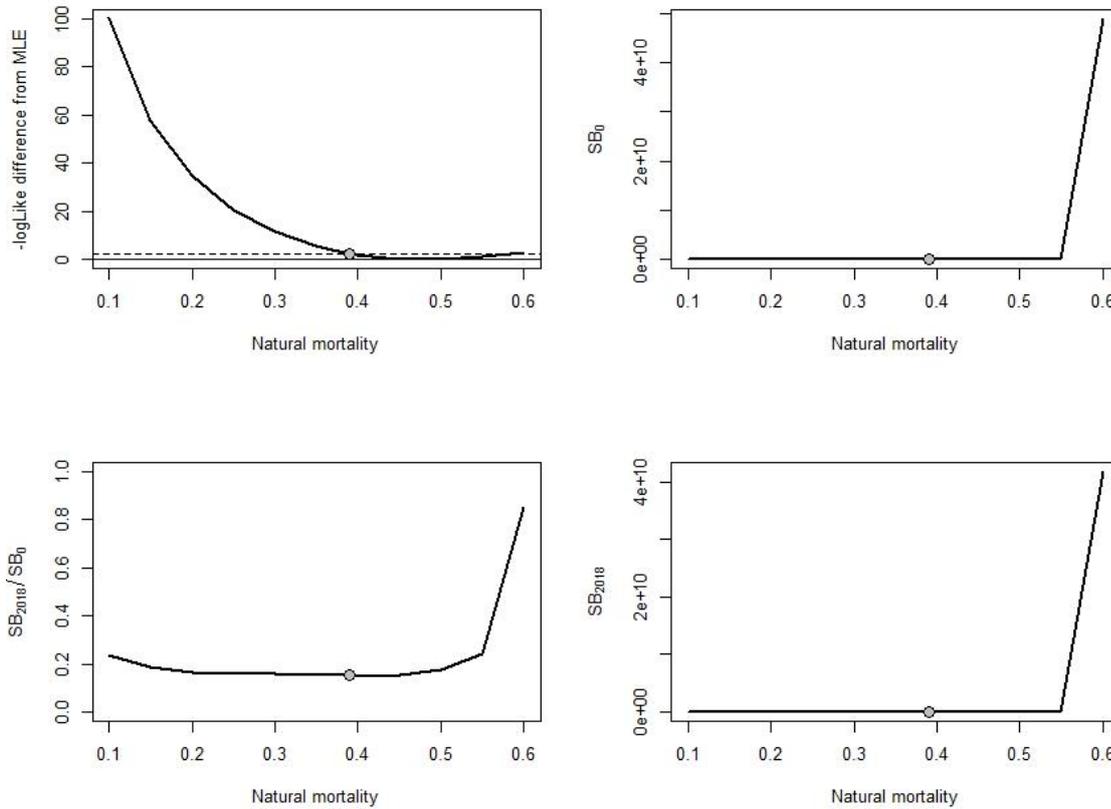


Figure D6: Likelihood profile in French Guiana red snapper SS model without the catch per unit effort index for natural mortality (M) and derived quantities (initial spawning output (SO_0); spawning output in 2018 (SO_{2018}), stock status (SO_{2018}/SO_0). The natural mortality of 0.39 estimated by “The Natural Mortality tool” is showed by a grey dot.

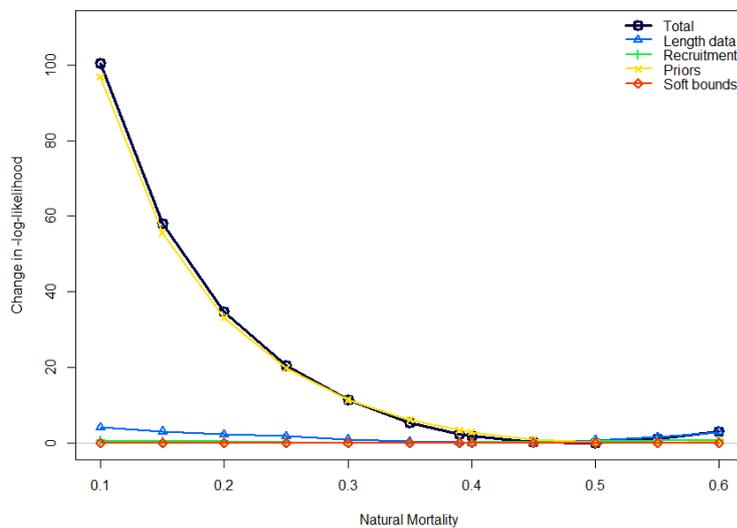


Figure D7: Likelihood component contributions for natural mortality (M) profile in French Guiana red snapper SS model without the catch per unit effort index.

Appendix E

Comparison of the length composition distribution and selectivity.

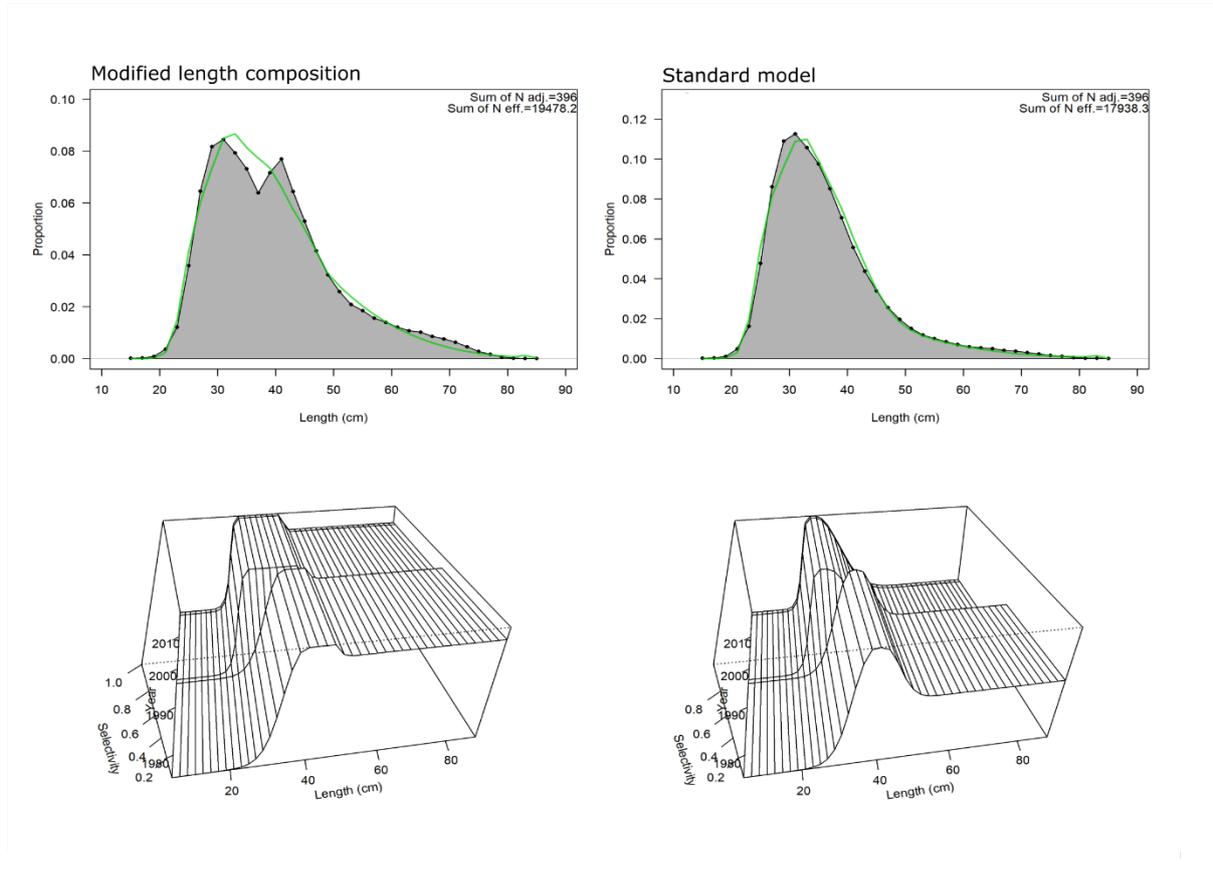


Figure E1: Comparison of the length composition distribution and selectivity for the reference model (right) and the model using the modified length composition dataset (left).

Appendix F

Comparison of the reference model and the model using a power link function.

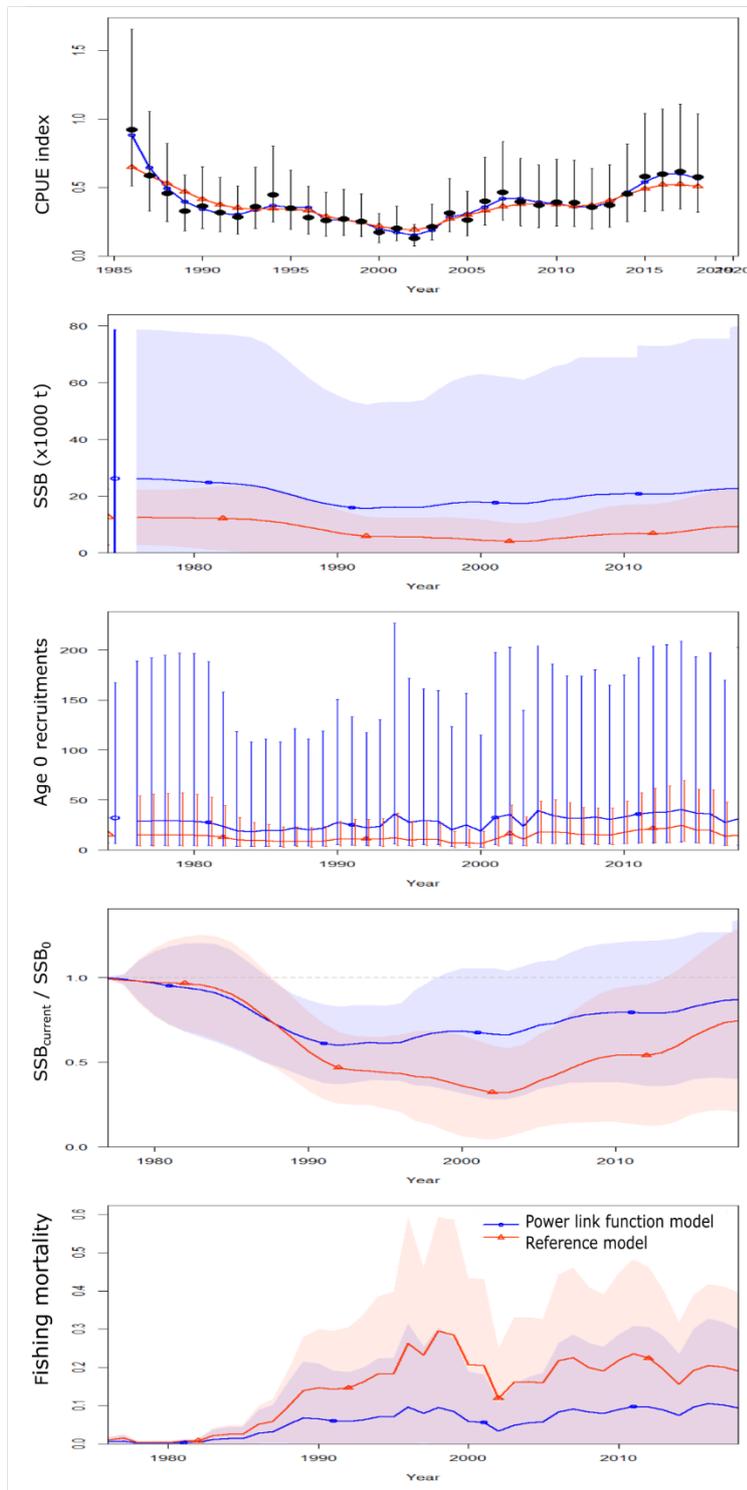


Figure F1: Comparison of the main model outputs and index fit for the reference model and the model using a power link function between the CPUE index and the catchability.